



Financial, institutional, and macroeconomic determinants of cross-country portfolio equity flows: The case of developed countries[☆]

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ABSTRACT

This paper examines the determinants of financial equity flows to investigate the role played by business cycles, government debt and sovereign rating scores, and whether the impact depends on the magnitude and direction of the flows. Using a new, richer dataset of flows among developed countries over 2001–2018, our key findings are as follows: (i) equity flows are more intense among countries at the same stage of the business cycle (ii) equity flows are higher to countries with a relatively lower debt to GDP ratio (iii) financial and macroeconomic variables are important for big equity flows, while institutional variables are important for the small flows. Overall, considering a wider range of factors under-explored in the literature, we provide a stronger understanding of the development of risks in the financial sector as well as the linkages with other sectors of the economy.

1. Introduction

The importance of cross-country capital flows is well understood in a world where the search for funding and yields tends to drive investments. In the 1990s, modeling exercises tended to adopt push/pull frameworks (Koepke, 2019; Levy Yeyati and Zúñiga, 2015). Subsequently, research also reflected the fact that financial equity flows could be impacted by economic shocks such as the global financial crisis (GFC), which created a period of extreme stress in the global financial markets and banking systems between mid-2007 and early 2009 (Fratzscher, 2012). Moreover, gravity-style variables supplemented the traditional push/pull framework (Araujo et al., 2017; Everett and Galstyan, 2020; Obstfeld and Rogoff, 2000; Portes et al., 2001; Portes and Rey, 2005). Despite this large body of research, there are important gaps. The impact of business cycle synchronization and

sovereign rating scores on bilateral equity flows is poorly understood. Moreover, the recent work of Wisniewski and Jackson (2021) suggests a negative relationship between the government debt-to-GDP ratio and stock market returns. However, further evidence is needed to confirm this relationship. Moreover, while research has stressed that different types of capital flows are driven by different sets of factors (Brafu-Insaidoo and Biekpe, 2014; Ibarra and Tellez-Leon, 2020), we have a limited understanding of the impact on the magnitudes of the flows. Providing a deeper examination of the determinants of financial flows is crucial so countries can design a mixture of policies to manage these flows.

Studying portfolio equity flows requires tackling significant data issues as discussed in Koepke and Paetzold (2020). In this present paper, we exploit new data for 40 economies, where each is treated

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as a reporter and partner over the 2001–2018 period. This provides us with 24,282 observations. This is a considerably larger sample than other studies in the research area; for example, [Kemme et al. \(2021\)](#) explored the determinants of equity flows with data covering 149 source countries and 34 OECD host countries, which provided 15,697 observations. Therefore, our data can be considered as a description of the population of all financial equity flows among developed countries from 2001 to 2018. With this data, we explore the features of cross-country equity flows among developed countries over the last two decades.

Our analysis of the entire sample provides baseline results that indicate that the role of financial factors is significant. On the one hand, we find and report that financial equity flows follow the highest return. On the other hand, financial equity flows exhibit no pattern of risk-sharing, with flows pursuing higher standard deviations of returns. The lack of risk-sharing is further heightened by flows among countries with highly correlated stock markets and experiencing the same phase of the business cycle. This result underpins the vast body of research that reports a lack of international risk-sharing as predicted by international business cycle models ([Mace, 1991](#); [Kose et al., 2009](#); [Pierucci and Ventura, 2010](#); [Lewis and Liu, 2015](#); [Fuleky et al., 2018](#)). We also find evidence supporting the role of the overall current account position, as well as bilateral imports and exports. However, The importance of a country's indebtedness is conditional on the sovereign rating status of the economy. We also find that institutional factors play a significant role in driving bilateral portfolio equity flows.

Nonlinearities present in the examined data motivated us to resort to quantile regression. In this analysis, we find that the importance of broadly defined groups of determinants is conditional on the size of the underlying flows. Financial factors, current account position, and degree of business cycle synchronization play a major role for the inflows and outflows of the highest magnitudes. Moreover, at the top and bottom quantiles of the flow sizes, we find that risk-sharing behavior can be observed, while there is no evidence of risk-sharing behavior in the middle quantiles. The middle quantiles of the bilateral flows are dominated by institutional factors: capital controls, political stability, and availability of information. We also see that in contrast to the ([Lucas, 1990](#)) paradox; portfolio equity flows from richer to poorer countries. Moreover, the results are significant only for inflows, indicating that capital not only flows to poorer countries but stays there.

Therefore, the contribution of this paper is three-fold. Firstly, we adopt a richer dataset that allows us to examine bilateral equity flows across developed countries over the last two decades. Secondly, we examine the importance of both countries being at the same or different stages of the business cycle, as well as differences in government debt and their sovereign rating scores. Finally, we consider the effects of the examined determinants on different magnitudes and directions of portfolio equity flows. This, in turn, enables us to demonstrate when a given group of factors has the most profound impact on bilateral capital flows. Therefore, in this paper, we answer several closely related research questions: (i) How do bilateral equity flows depend on factors that can be assigned to one of three broad driver categories: financial, macroeconomics, and institutional? Additionally, we assess whether (ii) equity flows are determined by the differences in government debt across countries, (iii) sovereign ratings are important when we seek to explain the drivers for bilateral equity flows, (iv) the factors affecting portfolio equity flows are related to their magnitude; (v) there is a distinction between the size of international equity flows and the category of determinants that primarily drives them.

The remainder of this paper is laid out as follows. Section 2 reviews the related literature, and Section 3 discusses the methodology, estimation, and data. The empirical results are discussed in Section 4. This includes the results for the main sample, sub-samples as well as semi-parametric and quantile regressions. Section 5 concludes.

2. Literature review

The early 20th and early 21st centuries are periods marked by very different trends in capital flows. [Schularick \(2006\)](#) show that flows from developed to developing countries were increasingly important in the early 20th century, but this is not the case in the 21st century as capital flows from developed to developing countries flattened out. Capital flows have also been the focus of concern when considering the risks associated with exchange rate fluctuations, capital that moves quickly and frequently (for further discussion on 'hot money' see [Yan \(2018\)](#) and on detecting surges in flows see [Kaya et al. \(2022\)](#)) as well as the loss of monetary control ([Binici et al., 2010](#)). These issues have led to a literature that explores the determinants of equity flows, where global/external (push) and country-specific (pull) factors are used to categorize the independent variables used in modeling exercises (see [Koepeke \(2019\)](#), [Levy Yeyati and Zúñiga \(2015\)](#) for reviews of the equity flows literature).

Typically, global factors include the general level of risk (negative relationship expected), interest rates (negative relationship expected) and international productivity levels (positive relationship expected), where the reference country group tends to be large developed countries. Promoted by various crises, push factors began to attract more interest in the 1990s. On the other hand, pull factors such as country-specific risks (negative relationship expected), rates of return (positive relationship expected) and productivity (positive relationship expected) were the focus of studies before the 1990s. During the GFC, this framework was also supplemented by 'shock' factors. Since the emergence of the push-pull framework in the 1990s, researchers began to consider the relative importance of these types of factors; [Fernandez-Arias \(1996\)](#) provided an early contribution to this literature, where they concluded that global factors were more dominant. More recently, [Sarno et al. \(2016\)](#) examined flows from the US to another 55 destinations and concluded that global factors appear more important than country-specific factors in explaining flows. [Fratzscher \(2012\)](#) also found that global factors were generally of the most importance during the financial crisis. Moreover, [Mandalinci and Mumtaz \(2019\)](#) find support for the push-pull framework and conclude that regional variations are more important than global variations in explaining portfolio capital flows to emerging economies. Indeed, several other studies report and stress the importance of global risk and liquidity constraints ([Forbes and Warnock, 2012](#); [Belke and Volz, 2019](#)), notably for emerging economies, while the potential role of geopolitical developments can also be a relevant determinant (see [Feng et al. \(2023\)](#)).

While the traditional push-pull framework remains popular, these factors have also increasingly been complemented by gravity-style variables and other variables that cannot easily be categorized into push/pull (e.g. contagion effects). [Araujo et al. \(2017\)](#), [Everett and Galstyan \(2020\)](#), [Obstfeld and Rogoff \(2000\)](#), [Portes et al. \(2001\)](#) and [Portes and Rey \(2005\)](#) have demonstrated that the gravity model can explain financial flows as well as trade flows. As an example, a typical gravity variable is distance, whereas in a trade context, this proxies for trade costs. In the context of capital flows, greater distance suggests less market information on which to base investment decisions. This has prompted the international capital flow literature to consider a range of institutional variables ([Lothian, 2006](#); [Montiel and Reinhart, 1999](#); [Neumann et al., 2009](#)). For instance, [Ftiti et al. \(2024\)](#), for the period 1995–2022, for 12 countries, report that the BRICS' inflow dynamics are more linked to domestic factors, and outflows are better linked to global determinants, and the opposite occurs for Europe.

Moreover, [Yang et al. \(2019\)](#) also reported that in emerging countries, capital flows increase following financial liberalization, but the magnitude of such flows decreases in that context for developed countries. Indeed, financial liberalization can be beneficial depending on the country's institutional setup (see [Chinn and Ito \(2006\)](#)) and provide access to a bigger pool of foreign capital reduce adverse selection and

moral hazard (see, [Mishkin et al. \(2003\)](#)). Nevertheless, capital controls might be useful in the context of financial crises, notably in the case of emerging market economies.

Therefore, in this paper, we assess the drivers of equity flows, using a range of push and pull variables in three main categories as controls. Indeed, [Calvo et al. \(1993, 1996\)](#) mentioned the relevance of pull (domestic) and push (global) factors as determinants of international capital flows. Hence, we control notably for: (i) financial factors with our variables related to stock market indices and also split our sample according to sovereign rating notations; (ii) institutional country specific factors are accounted for via measures of capital controls and institutional quality; (iii) macroeconomic factors are controlled for by variables for trade, public debt, exchange rates ([Brooks et al. \(2004\)](#), reported a strong relationship between exchange rate developments and equity flows) and GDP. We have two avenues of particular interest. Firstly, we examine whether the influence of a specific determinant of portfolio equity flows is conditional on the magnitude and direction of the flows using a quantile regression approach. Secondly, we focus our attention on novel factors, such as business cycle synchronization, government debt and sovereign rating scores. Both aspects are under-explored in the literature.

There is research considering the impact of European Central Bank monetary policies on stock markets ([Haitsma et al., 2016](#)) as well as the impact of fiscal and monetary shocks on stock market performance ([Afonso and Sousa, 2011](#); [Chatziantoniou et al., 2013](#)). Furthermore, there is also some evidence of a link between capital flows and global business cycles ([Kose et al., 2008, 2012](#); [Eller et al., 2020](#)). However, to the best of our knowledge, there is very limited consideration of the impact of business cycle synchronization on bilateral equity flows. We also specifically consider whether there are different effects depending on sovereign rating scores. Similarly, there is limited research considering this dimension to explain cross-country bilateral equity flows; for exceptions, see [Kim and Wu \(2008\)](#) who consider the long-term effect of sovereign rating scores on equity flows to emerging economies, and [Christopher et al. \(2012\)](#) who explore the connection between regional stock market co-movements and sovereign rating scores for emerging economies. There is also only a limited amount of research examining the related issue of the impact of sovereign rating scores on international banking flows ([Kim and Wu, 2008](#)). Thirdly, the recent work of [Wisniewski and Jackson \(2021\)](#) suggests a negative relationship between the government debt-to-GDP ratio and stock market returns. This is also a very under-explored area, with older contributions, such as the research based on Canadian data by [Darrat \(1990\)](#).

3. Methodology

3.1. Data and variables under investigation

The dataset used to construct the dependent variable comes from Finflows database ([Nardo et al., 2017](#)). This research utilizes data on annual portfolio equity inflows over the period 2001–2018 among the following 40 economies¹: Austria, Australia, Belgium, Canada, Switzerland, Chile, Czechia, Germany, Estonia, Spain, Finland, France, Greece, Hong Kong, Hungary, Ireland, Iceland, Italy, Japan, Korea, Lebanon, Lithuania, Luxembourg, Latvia, Malta, Mauritius, Mexico, Netherlands, Norway, New Zealand, Panama, Poland, Portugal, Romania, Sweden, Singapore, Slovakia, Turkey, the UK, and the USA. With 40 countries, there is a total of 1560 pairs of countries, where each country is treated as a reporting country and a partner country. However, due to missing observations, in this research we utilize the data on 1349 country pairs. Therefore, the total number of observations amounts to $18 \times 1349 =$

24,282. The list of missing country pairs is displayed in [Appendix A](#). In this setting, variable $PEinflows_{ijt}$ is defined as portfolio equity inflow to country i from country j in year t . Moreover, we explore potential determinants of portfolio equity inflows that can be divided into the following three main categories: financial factors, institutional factors, and macroeconomic factors. Summary statistics on all the examined variables are depicted in [Tables 7 and 8 in Appendix G](#).

3.1.1. Financial factors

The main drivers of equity flows are associated with financial factors. The empirical specification we use follows from the CAPM model ([Sharpe, 1964](#); [Lintner, 1965](#); [Mossin, 1966](#)) that was further developed into ICAMP — its international counterpart ([Solnik, 1974](#); [Merton, 1980](#); [Adler and Dumas, 1983](#); [Lewis, 1999](#)). Within this theoretical framework, portfolio allocation depends on three main factors expected returns (proxied by mean returns), expected volatility of the returns (proxied by standard deviations of the returns), and covariances between the returns on different assets (proxied by correlation coefficients). However, the model assumes frictionless trade in assets, as well as ignoring the potential impact of other factors that have been regarded as important in the existing literature. This point describes the three main financial determinants associated with ICAMP, while points 3.1.2 and 3.1.3 describe institutional and macroeconomic factors that were added to the basic model in the subsequent research.

The three main financial variables we examine as potential determinants of international equity flows are means, standard deviations, and correlation of returns on stock indices. In fact, we can consider the domestic returns and volatility as push factors and foreign returns and volatility as pull factors. Additionally, covariances or correlations of returns can be perceived as pull factors. To calculate the measures, we obtained monthly data on the values of major stock indices expressed in US dollars² in the examined countries for the 2000–2018 period. The list of all stock market indices used in the analysis can be found in [Appendix B](#). Data on stock market returns comes from the Thompson Reuters database. As the data on current values of returns is not known to the investors, we are using the lagged values in the research.³ Utilization of the lagged values also helps resolve the endogeneity issues, as current flows might influence the value of the returns. Within this setting, we constructed three financial variables. The difference in mean returns is defined as:

$$Rdif_{ijt} = MR_{it} - MR_{jt} \quad (1)$$

where: MR_{it} and MR_{jt} are mean monthly returns calculated over the 12 month period, between stock indices in country i and country j , respectively, in year t .⁴ The difference in standard deviations is calculated as:

$$SDdif_{ijt} = SDR_{it} - SDR_{jt} \quad (2)$$

where: SDR_{it} and SDR_{jt} are standard deviations of monthly returns calculated over the 12 month period, between stock indices in country i and country j , respectively, in year t .⁵ Finally, Cor_{ijt} denotes the correlation coefficient of monthly returns calculated over the 12 month period, between stock indices in country i and country j , respectively, in year t . The percentiles of the dependent variable, as well as the financial variables, are depicted in [Figs. 1 and 2](#).

² We ran robustness checks using values expressed in local currency. The results are available in [Appendix C, D, and E](#)

³ Using contemporary values produces quantitatively similar results as reported here.

⁴ We obtained qualitatively similar results using a logarithmic specification: $\ln[(MR_{it} + 1)/(MR_{jt} + 1)]$.

⁵ We obtained qualitatively similar results using a logarithmic specification: $\ln[(SDR_{it})/(SDR_{jt})]$.

¹ The group that, using the nomenclature from The Economist, could be called “mostly developed economies”.

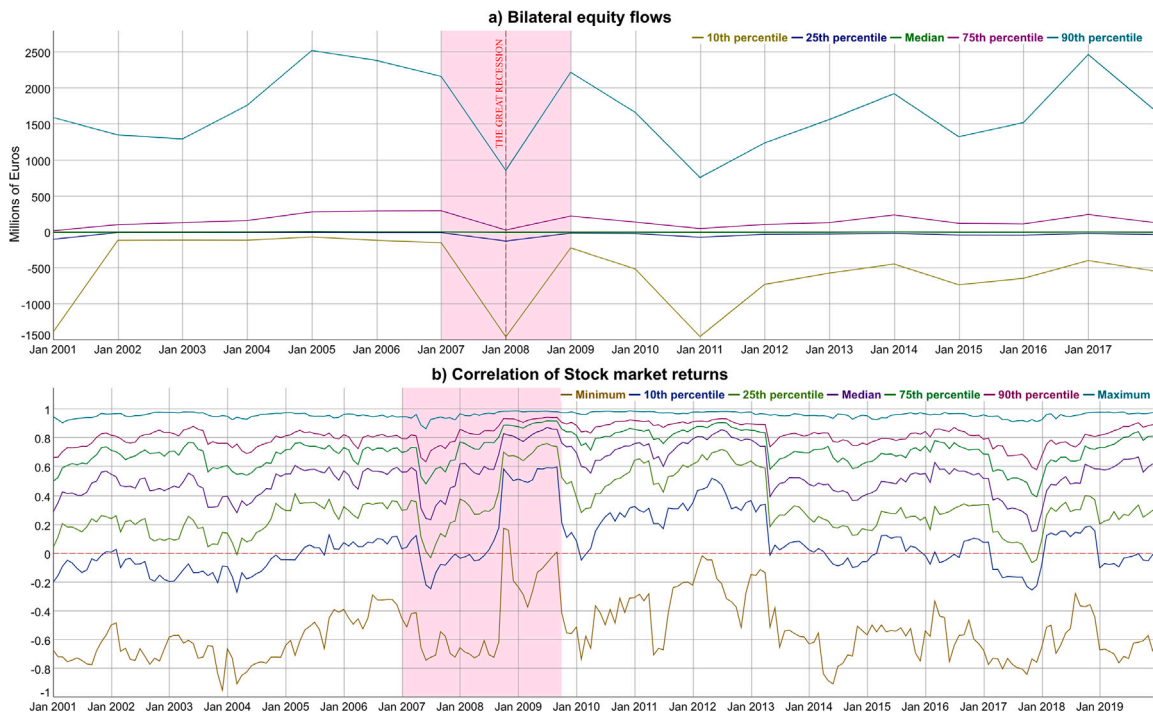


Fig. 1. Percentiles of the distribution of portfolio equity flows and correlations of returns over time.

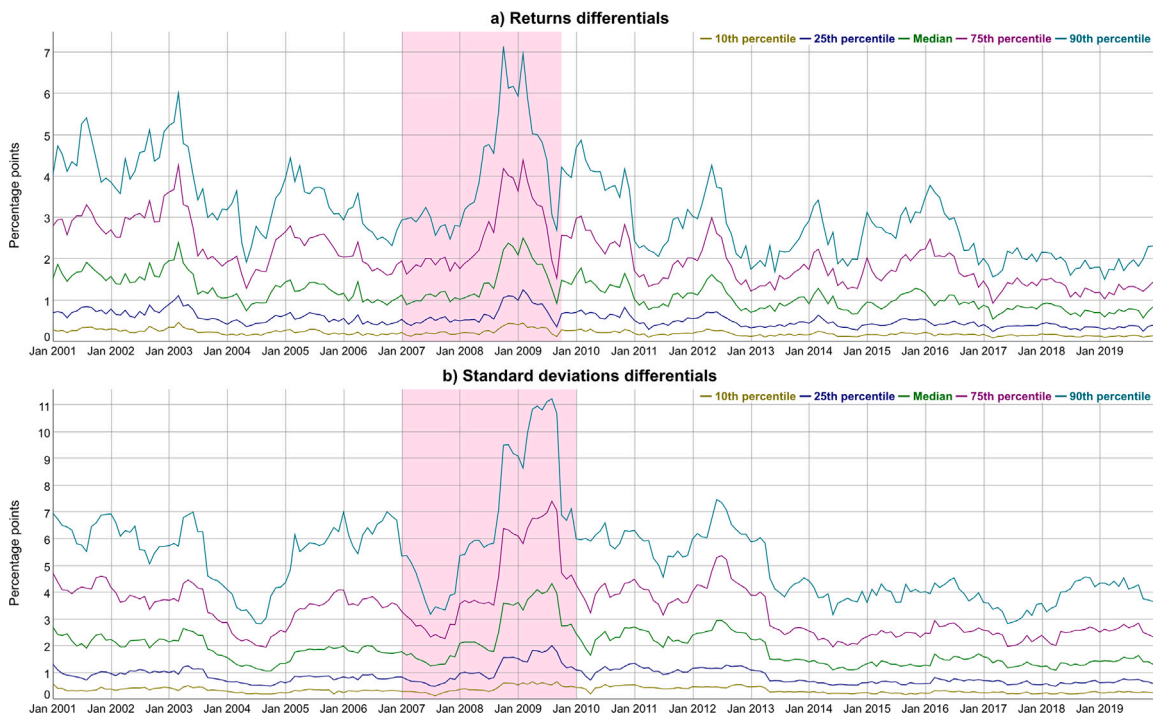


Fig. 2. Percentiles of the distribution of mean returns and standard deviations over time.

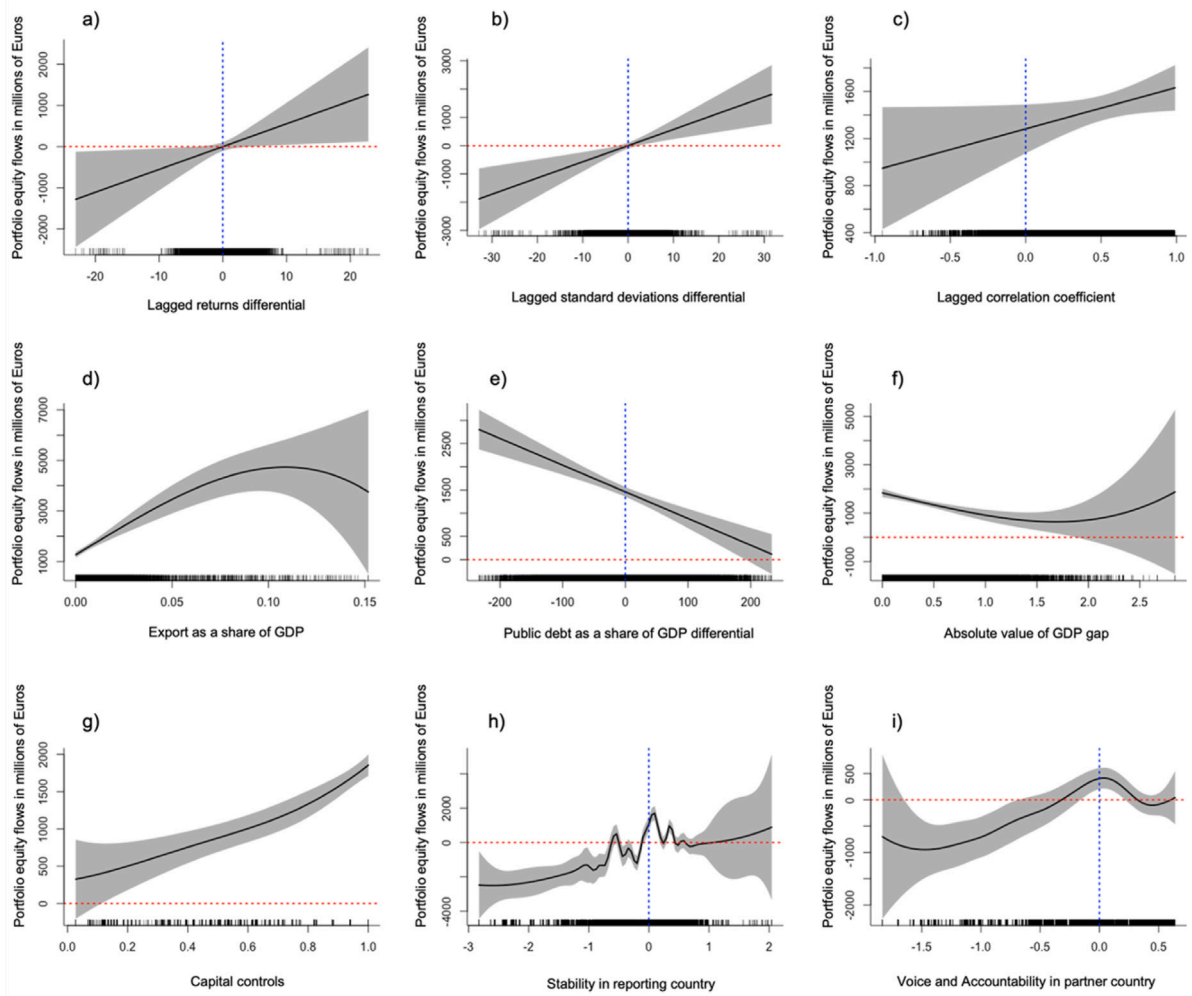


Fig. 3. Results of semi-parametric regression.

3.1.2. Institutional factors

The first institutional factor is the degree of capital controls across examined countries. To construct this variable, we utilized the (Chinn and Ito, 2006) database on de jure measures of financial openness. The measure of financial openness in this database (FO_{it}) for a given country i takes values from 0 (indicating no capital mobility) to 1 (indicating perfect capital mobility). As the capital mobility across pairs of countries depends on the degree of controls in both countries, we define the bilateral measure of capital controls as:

$$CapControls_{s_{ijt}} = FO_{it} * FO_{jt} \quad (3)$$

The advantage of using a product lies in the fact that the measure is bound between 0 and 1 and can take the value of 0, even if one of the countries is characterized by perfect capital mobility, while the other imposes prohibitive capital controls. Indeed, when it comes to institutional factors, we mostly have barriers and deterrents, so they are considered either push or pull factors. In this case, and for a specific country, we can study a pull factor and a push factor, given the foreign country's degree of capital mobility. The role of capital market controls in impeding capital flows has been examined by Montiel and Reinhart (1999), Kalemli-Ozcan et al. (2010), Pasricha et al. (2018), Boero et al.

(2019), Frost et al. (2020), Mercado and Noviantie (2020), Bricongne et al. (2021), and Mercado (2023a).

For the construction of another four measures of institutional quality, we utilized the World Bank Worldwide Governance Indicators database. In order to construct these measures, we used two indicators available in this database. The first of them is the Voice and Accountability measure, which, besides freedom of expression and freedom of association, captures the availability of information to citizens, which can be crucial for making informed decisions about international investments. The values of these measures for reporter and partner country, $RapVaA_{it}$ and $ParVaA_{jt}$, respectively, serve as proxies for the availability of information. The link between information availability and international financial flows was examined by Portes et al. (2001), Portes and Rey (2005), Houston et al. (2012), and Choi et al. (2014). The second variable from the database we use is Stability and Absence of Violence, which represents the stability of the political system and proxies the probability that the investment can be appropriated by the new government. We construct the measure for both reporter and partner country, $RapStab_{it}$ and $ParStab_{jt}$, respectively. The role of the quality of institutions in facilitating cross-border financial flows was examined by Papaioannou (2009), Houston et al. (2012), Forbes

et al. (2015), Bricongne et al. (2021), Janus (2023), Mercado (2023a), and Mercado (2023b).

3.1.3. Macroeconomic factors

The third group of examined variables we consider are macroeconomic variables. Firstly, we examined the impact of bilateral imports and exports expressed as a share of the reporting country's GDP, $Import_{ijt}$ and $Export_{ijt}$, respectively. The role of trade in facilitating cross-border financial flows was examined by Portes and Rey (2005), Binici et al. (2010), Ghosh et al. (2014), Hobza and Zeugner (2014), and Mercado (2023b). We also examine the impact of the current account position of the reporting country expressed as a share of GDP, CA_{it} . The impact of the current account position on the size of international financial flows was considered by De Gregorio et al. (2000), Houston et al. (2012), Ghosh et al. (2014), and Janus (2023). In addition, we explore the role of government debt by calculating the following variable:

$$DBdif_{ijt} = Debt_{it} - Debt_{jt} \quad (4)$$

where $Debt_{it}$ and $Debt_{jt}$ are the debt-to-GDP ratios in country i and country j , respectively, in year t . The examination of the impact of public finance on the size and direction of financial flows was undertaken by Herrmann and Mihaljek (2013) and Janus (2023). To establish the impact of exchange rate volatility on portfolio equity flows, we used the data on monthly bilateral nominal exchange rates. Then, we calculated the measure of exchange rate volatility as:

$$Exchange_{ijt} = \frac{SD(BiER_{ijt})}{M(BiER_{ijt})} \quad (5)$$

where: SD and M denote, standard deviation and mean, while $BiER_{ijt}$ is a series of monthly bilateral exchange rates across country i and country j , in year t . The division of the standard deviation by the mean has the advantage of expressing the volatility as a percentage deviation from the mean, thus facilitating better comparisons across pairs of countries with high and low absolute levels of bilateral exchange rates. The exchange rate volatility has been considered an impediment to financial flows by Jonson et al. (1982), Herrmann and Mihaljek (2013), Forbes et al. (2015), Gelman et al. (2015), and Lu et al. (2022).

Moreover, we examine the role of the difference in the level of development using the difference in the level of GDP per capita. The measure is calculated as:

$$GDP_{pc}dif_{ijt} = GDP_{pcit} - GDP_{pcjt} \quad (6)$$

where GDP_{pcit} and GDP_{pcjt} is GDP per capita of country i and country j , respectively, in year t .⁶ The impact of the difference in the level of development has been examined by Ghosh et al. (2014), Houston et al. (2012), Forbes et al. (2015), Mercado (2023a).

In order to establish the role of business cycle synchronization in determining the size of portfolio flows, we first collected that data on real GDP and used the Hodrick–Prescott filter to calculate the output gaps. The dummy variables BCS_{ijt} take the value of 1, when both countries i and j have positive or negative output gaps in year t , and 0 otherwise. In the case of the semi-parametric regressions, where utilization of the binary variable is inappropriate, we used a measure of business cycle co-movement defined as:

$$GAPdif_{ijt} = GAP_{it} - GAP_{jt} \quad (7)$$

⁶ We obtained qualitatively similar results using a logarithmic specification: $\ln[(GDP_{pcit})/(GDP_{pcjt})]$.

where GAP_{it} and GAP_{jt} are the output gaps in country i and country j , respectively, in year t .

Finally, to control for the sizes of the examined economies, we use the product of GDPs in country i and country j , respectively, in year t , $GDPprod_{ijt}$. The impact of the market sizes, or “the gravity effect” was examined by Mercado and Noviantie (2020), Bricongne et al. (2021), Mercado (2023a) and Mercado (2023b). The data on macroeconomic variables comes from the IMF Directions of Trade, IMF World Economic Outlook, IMF International Financial Statistics, and Penn World Table.

3.2. Estimation strategy

Regarding our estimation strategy, we estimate the following equation as our baseline:

$$\begin{aligned} PEinflows_{ijt} = & \beta_1 Rdif_{ijt-1} + \beta_2 SDdif_{ijt-1} + \beta_3 Cor_{ijt-1} \\ & + \beta_4 RepStab_{it} + \beta_5 ParStab_{jt} + \beta_6 RepVaA_{it} + \beta_7 ParVaA_{jt} \\ & + \beta_8 CapControls_{ijt} + \beta_9 Import_{ijt} + \beta_{10} Export_{ijt} + \beta_{11} CA_{it} \\ & + \beta_{12} DBdif_{ijt} + \beta_{13} BCS_{ijt} + \beta_{14} Exchange_{ijt} \\ & + \beta_{15} GDP_{pc}dif_{ijt} + \beta_{16} GDPprod_{ijt} + \eta_{ij} + \zeta_t + \varepsilon_{ijt} \end{aligned} \quad (8)$$

where the abbreviations of all the variables were explained in the previous subsection, η_{ij} is the country-pair specific fixed effect, ζ_t is the time effect, and ε_{ijt} denotes the stochastic component. As mentioned earlier, we are using lagged values for $Rdif$, $SDdif$, and Cor , as information about the contemporary values of mean returns, standard deviations, and correlations is unavailable to the agents making the trade. Additionally, using the lagged variables solves the problem of endogeneity between the portfolio flows and the aforementioned variables.

We proceed with the estimations in three steps. Firstly, we estimate only the financial equation:

$$PEinflows_{ijt} = \beta_1 Rdif_{ijt-1} + \beta_2 SDdif_{ijt-1} + \beta_3 Cor_{ijt-1} + \eta_{ij} + \zeta_t + \varepsilon_{ijt} \quad (9)$$

denoted as Model 1. The next two considered variants add all macroeconomic variables, however they differ in the use of the institutional variables. In the first variant, Model 2, we use only variables associated with political stability:

$$\begin{aligned} PEinflows_{ijt} = & \beta_1 Rdif_{ijt-1} + \beta_2 SDdif_{ijt-1} + \beta_3 Cor_{ijt-1} \\ & + \beta_4 RepStab_{it} + \beta_5 ParStab_{jt} \\ & + \beta_8 CapControls_{ijt} + \beta_9 Import_{ijt} + \beta_{10} Export_{ijt} + \beta_{11} CA_{it} \\ & + \beta_{12} DBdif_{ijt} + \beta_{13} BCS_{ijt} + \beta_{14} Exchange_{ijt} \\ & + \beta_{15} GDP_{pc}dif_{ijt} + \beta_{16} GDPprod_{ijt} + \eta_{ij} + \zeta_t + \varepsilon_{ijt} \end{aligned} \quad (10)$$

while in the second variant, Model 3, we consider institutional variables associated with availability of information:

$$\begin{aligned} PEinflows_{ijt} = & \beta_1 Rdif_{ijt-1} + \beta_2 SDdif_{ijt-1} + \beta_3 Cor_{ijt-1} \\ & + \beta_6 RepVaA_{it} + \beta_7 ParVaA_{jt} \\ & + \beta_8 CapControls_{ijt} + \beta_9 Import_{ijt} + \beta_{10} Export_{ijt} + \beta_{11} CA_{it} \\ & + \beta_{12} DBdif_{ijt} + \beta_{13} BCS_{ijt} + \beta_{14} Exchange_{ijt} \\ & + \beta_{15} GDP_{pc}dif_{ijt} + \beta_{16} GDPprod_{ijt} + \eta_{ij} + \zeta_t + \varepsilon_{ijt}. \end{aligned} \quad (11)$$

⁷ To the best of our knowledge, the impact of synchronization of business cycles have not been researched thus far. This gap in the literature is particularly surprising as it is one of the main features of international business cycle literature Backus et al., 1992; Backus and Smith, 1993. However, the impact of financial flows on the degree of business cycle synchronization has been examined by Beck (2021a,b).

Table 1
Sovereign ratings.
Source: The authors.

Characterization of debt and issuer (source: Moody's)	Rating			
	S&P	Moody's	Fitch	Scale
Highest quality	AAA	Aaa	AAA	17
	AA+	Aa1	AA+	16
High quality	AA	Aa2	AA	15
	AA-	Aa3	AA-	14
Strong payment capacity	Investment grade			
	A+	A1	A+	13
	A	A2	A	12
	A-	A3	A-	11
Adequate payment capacity	BBB+	Baa1	BBB+	10
	BBB	Baa2	BBB	9
	BBB-	Baa3	BBB-	8
Likely to fulfill obligations, ongoing uncertainty	Speculative grade			
	BB+	Ba1	BB+	7
	BB	Ba2	BB	6
	BB-	Ba3	BB-	5
High credit risk	B+	B1	B+	4
	B	B2	B	3
	B-	B3	B-	2
	CCC+	Caa1	CCC+	
Very high credit risk	CCC	Caa2	CCC	
	CCC-	Caa3	CCC-	
	CC	Ca	CC	1
Near default with possibility of recovery			C	
			DDD	
Default	SD	C	DD	
	D		DD	
			D	

Finally, the last specification, Model 4, considers only those institutional variables that were statistically significant in Model 2 or Model 3. In the main results, Model 4 takes the form:

$$\begin{aligned}
 PEinflows_{ijt} = & \beta_1 Rdif_{ijt-1} + \beta_2 SDDif_{ijt-1} + \beta_3 Cor_{ijt-1} \\
 & + \beta_4 RepStab_{it} + \beta_7 ParVaA_{jt} \\
 & + \beta_8 CapControls_{ijt} + \beta_9 Import_{ijt} + \beta_{10} Export_{ijt} + \beta_{11} CA_{it} \\
 & + \beta_{12} DBDif_{ijt} + \beta_{13} BCS_{ijt} + \beta_{14} Exchange_{ijt} \\
 & + \beta_{15} GDP_{pc} dif_{ijt} + \beta_{16} GDP_{prod} _{ijt} + \eta_{ij} + \zeta_t + \epsilon_{ijt}.
 \end{aligned}
 \tag{12}$$

However, in the sub-samples, Model 4 is specified differently, depending on the results obtained for Models 2 and 3.

Furthermore, we examine whether the results obtained using the entire sample hold up in several sub-samples. Firstly, we consider splitting the sample between pairs of countries associated with different sovereign rating categories. In Table 1, and following Afonso et al. (2014), we explain the quantitative rating scale, from 1 (lowest quality) to 17 (highest quality, AAA), used to categorize the respective qualitative ratings from the three main rating agencies (Moody's, S&P and Fitch).

Hence, the sample is divided into pairs characterized by AAA rating alone, pairs with a rating below AAA, and pairs where one of the countries has an above AAA rating while the other one is below AAA. Secondly, in Appendix F, we present the results of the split based on investment grade rating, i.e., BBB or higher. In this case, the sample is divided into a sample of pairs with both countries characterized by investment grade, pairs with both countries below investment grade, and pairs where one country has investment grade and the other country has not.

Finally, we divide the sample into two consecutive sub-periods: from 2001 to 2009 and from 2010 to 2018, which accounts for the potential relevance of the GFC.

To examine possible nonlinearities in the way the determinants influence portfolio equity flows, we have used a semi-parametric regression approach (Ruppert et al., 2003). Consequently, we estimated

equations of the form:

$$PEinflows_{ijt} = f(Y_{ijt}) + \sum_{q=1}^{15} \delta_q X_{ijt}^q + \eta_{ij} + \zeta_t + \vartheta_{ijt} \tag{13}$$

where Y_{ijt} denotes the observation of a chosen variable from the 16 described above, X_{ijt}^q is one of the 15 remaining variables, indexed by q , used as linear controls. η_{ij} is the country-pair specific fixed effect, ζ_t is the time effect, and ϑ_{ijt} denotes the stochastic component. $f(\cdot)$ denotes a function fitted using radial basis functions (French et al., 2001), which is a generalization of the penalized spline smoother (Eilers and Marx, 1996; Ruppert and Carroll, 2000). The smoothing parameters selection is performed using restricted maximum likelihood, and $\hat{f}(Y_{ijt})$ is obtained with estimated best linear unbiased prediction (Robinson, 1991).

We have estimated our main equation (12), resorting to a quantile regression. The main advantage of a quantile regression approach relies on the analysis of the relationships of explained and explanatory variables outside the average values of the data, allowing, at the same time, for analyzing possible non-linear relationships between the set of explanatory factors and the variable of interest. Consequently, the purpose of resorting to this methodology is to disclose the heterogeneous impacts of financial, institutional and macroeconomic variables over PEflows. Therefore, we divided our sample into ten quantiles, from the highest portfolio equity outflows (negative PEflows) to the highest portfolio equity inflows (positive PEflows), where this variable is a function of the above-mentioned financial, institutional and macroeconomic factors.

4. Empirical results

4.1. Main results

The main results from the full sample are shown in Table 2. Starting with the financial variables, the estimates suggest that equities are purchased in countries with higher rates of return along the lines predicted by the classical Markovitz model. However, against the prediction of the Markovitz portfolio analysis model, the money flows to countries with higher standard deviations of rates of return and across countries with highly correlated rates of return. Consequently, we do not see a behavior that could be described as risk-sharing, on the contrary, we see behavior that could be described as risk-seeking, and where investors tend to “hunt for yield” and chase investments with higher yields. In fact, the flows of financial resources among economies registering highly correlated rates of return are explained by the “trend-chasing” hypothesis. As explained in Kanas and Karkalakos (2017), higher return differentials act as a positive stimulus on financial flows to the economy registering higher returns. Therefore, the investor re-balances their financial portfolios, strengthening the correlation among the countries.

On the other hand, the recent literature has associated risk-taking behavior, translated as higher volatility in assets' returns and higher financial inflows, which can support our findings. For instance, Dinger and te Kaat (2020) show that the European banking sector plays a crucial role in explaining the positive correlation between risk and cross-country capital flows. As detailed by these authors, a large share of financial inflows are associated with higher-risk granted loans, evidencing agency problems, where this effect is reduced for smaller banks. Moreover, Tobe (2015) provides evidence that capital inflows lead to a pro-cyclical effect on asset price leverage with a corresponding increase in the overall risk.

In terms of macroeconomic variables, we see that the portfolio equity holdings by foreigners increase in countries with a current account deficit and in countries with a relatively lower debt-to-GDP ratio, highlighting the relevance of sounder fiscal policies for such investment decisions. Both results are in line with the standard international macroeconomics proposals. We also find intensified equity purchases

Table 2
Main results.

Variable	Model 1	Model 2	Model 3	Model 4
Rdif	6205** (2509)	6346** (2483)	5284** (2491)	5594** (2486)
SDdif	8772*** (1490)	9582*** (1549)	7729*** (1542)	8673*** (1545)
Cor	2066*** (174)	577*** (190)	491** (195)	419** (193)
CA		-1357*** (366)	-837** (361)	-1197*** (367)
CapControls		1537*** (196)	1483*** (202)	1209*** (199)
Exchange		3009 (1977)	3108 (1977)	3085 (1976)
RepStab		367*** (90)		400*** (90)
ParStab		2.34 (86)		
RepVoice			-16.52 (113)	
ParVoice			444.1*** (106.3)	483.3*** (106.2)
Export		28920*** (4781)	28260*** (4777)	28770*** (4767)
Import		23060*** (4435)	22740*** (4435)	21600*** (4440)
GDPpedif		-0.00*** (0.00)	-0.00*** (0.00)	-0.00*** (0.00)
BCS		333*** (109)	340.8*** (108.6)	348.3*** (108.6)
DebtDif		-9.90*** (2.97)	-9.77*** (2.96)	-9.96*** (2.96)
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)
Sample size	24282			

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimators.

across countries with larger trade flows, as both exports and imports, as shares of GDP, contribute positively to those flows. Moreover, equity flows are more intense among countries in the same phase of the business cycle. This reinforces the case against risk sharing taking place, which is also against predictions of international business cycle theory (see the discussion about the positive comovement between business cycles synchronization and international financial flows, and to what extent it can be positive for the European integration, in Beck (2020)). Another macroeconomic variable with significant results is the GDP per capita difference. Here we find that capital flows from richer to poorer countries — contrary to the Lucas paradox (Lucas, 1990). However, and based on the literature that devotes some attention to the Lucas paradox, there is a common feature that advances an explanation for this paradox: institutions. In fact, institutional quality tends to eliminate the so-called Lucas paradox (see the discussion of the effect of institutional quality on the validity of the Lucas paradox in Azémar and Desbordes (2013), Göktan (2015), Akhtaruzzaman et al. (2018)). Therefore, and given the institutional aspect that we also analyze, we cannot conclude that our result is entirely surprising, given the positive institutional effect in explaining migration flows.

In fact, when it comes to institutional variables, there are three results to highlight. Firstly, we observe higher capital flows across countries characterized by higher capital mobility. Secondly, we see higher inflows to countries characterized by higher political stability, and finally, we see higher inflows from countries characterized by higher “Voice and Accountability”, which proxies for the availability of information. Finally, the product of the GDPs of the two countries is always significant. This serves as a control for the size of the economies, as our variables of interest are total flows.

4.2. Results from sub-samples

4.2.1. Sovereign ratings

We are also interested in exploring whether sovereign ratings are important when we seek to explain the drivers for bilateral equity flows. Therefore, we divide our sample according to whether the bilateral equity flows are across countries both with a AAA rating, both with a below AAA rating, or one country with a AAA rating and the other a below AAA rating. While the main results suggest that equities are generally purchased in countries with higher rates of return, the sub-sample results in Table 3 indicate that it is in the country-pairs that both have a sovereign rating below AAA (Table 3, panel b) that tend to purchase equities in countries with a higher rate of return. There is little evidence for the relevance of the differences in returns for the sub-sample of countries with an AAA rating (panel a), or for the group where flows are across an AAA rated country and a below AAA rated country (panel c). The correlation of stock market indices remains significant in the case where bilateral equity flows are across countries both with a AAA rating (panel a) and both with a below AAA rating (panel b), but not in the case where one country has a AAA rating and the other has a below AAA rating (panel c). In addition, the difference in variances remains important in all sub-samples, again indicating risk seeking.

The first macroeconomic variable, *CA*, which is the current account position of the reporting country as a share of GDP, is not significant when bilateral equity flows are across countries both with a AAA rating or both with a below AAA rating. However, in the case where one country has a AAA rating and the other a below AAA rating (panel c), we find that *CA* is negative and significant, as was the case in the main results. Therefore, portfolio equity holdings increase in countries with a current account deficit when there is a difference in the sovereign rating of the two countries. Additionally, our earlier finding, from the baseline results, was that equity flows are more intense among countries in the same phase of the business cycle. However, the results by sub-sample suggest that this is only the case when equity flows are among countries both with a below AAA rating. Furthermore, we see that cross-country capital flows are explained by the existence of a relatively lower debt-to-GDP ratio only when we consider countries both with a below AAA rating, or one country with a AAA rating and the other a below AAA rating. The insignificant result for countries both with a AAA rating (Table 3, panel a) makes intuitive sense since, typically, such countries should depict a better fiscal position, a key feature for the rating agencies. Similarly intuitive results are found when referring to the difference in GDP per capita, which is significant only in the case where one country has a AAA rating and the other a below AAA rating.

Turning to institutional factors, we observe higher capital flows among countries characterized by higher capital mobility in the case of both countries having a below AAA rating (Table 3, panel b), or one country with a AAA rating and the other a below AAA rating (panel c). For countries both with a AAA rating, capital controls are usually equal to 1, meaning that there are no barriers to capital movements and therefore, an insignificant result is expected. We do not find a significant link between equity flows and political stability when both countries have a AAA rating. However, the result from our main findings, which was that higher inflows to countries characterized by higher political stability, remains valid in the case of the other two categories. Finally, we see higher inflows from countries characterized by higher “Voice and Accountability”, which proxies for the availability of information, is only important when one country has a AAA rating and the other a below AAA rating (Table 3, panel c).

Table 3
Results from sub-samples: sovereign ratings.

Subsample	Flows among AAA rating countries				Flows among below AAA rating				Flows among AAA and below AAA rating			
	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4
Rdif	22892 (15760)	22490 (15950)	28060* (15900)	22640 (15900)	6973** (3458)	6492* (3417)	5987* (3423)	6676* (3414)	3826 (3550)	3383 (3528)	1817 (3543)	2090 (3534)
SDdif	20113** (9448)	16620* (9851)	21300** (9796)	16710* (9823)	6524*** (2052)	6997*** (2127)	5639*** (2122)	7486*** (2099)	10964*** (2108)	10740*** (2221)	8905*** (2207)	9747*** (2215)
Cor	5751*** (920)	2539** (1230)	2415* (1263)	2536** (1230)	2225*** (260)	880*** (275)	903*** (279)	903*** (274)	1287*** (246)	68 (266)	-101 (271)	-146 (269)
CA		-1877 (1170)	-1740 (1167)	-1860 (1162)	1013 (669)	1659** (667)	961 (668)		-1997*** (475)	-1499*** (469)	-1750*** (476)	
CapControls		2064 (1324)	-167 (1257)	2136* (1197)	1498*** (279)	1786*** (291)	1609*** (267)		1618*** (273)	1433*** (277)	1232*** (276)	
Exchange		17390* (10550)	15990 (10570)	17430* (10540)	2064 (2770)	1793 (2774)	1982 (2769)		2771 (2797)	2853 (2793)	2614 (2793)	
RepStab		71.9 (566.8)			564*** (127)		555*** (127)		282.6** (129)			348*** (128)
ParStab		-2073*** (496)		-2068*** (494)	176 (124)				160 (124)			
RepVoice			319 (453)			-881 (184)				53 (151)		
RepVoice			484 (421)			239 (169.2)				670*** (142)		718*** (142)
Export		78650*** (12130)	72570*** (11950)	78410*** (11980)	17010* (10200)	17150* (10210)	17800* (10190)		13950** (6145)	13950** (6138)	14310** (6123)	
Import		21260* (11540)	32960*** (11210)	21360* (11510)	47880*** (9313)	49470*** (9320)	47690*** (9313)		11190** (5697)	9563* (5692)	8796 (5695)	
GDPpedif		-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)		-0.00** (0.00)	-0.00** (0.00)	-0.00** (0.00)	
BCS		54.6 (424)	7.1 (425)	56.5 (423)	607.9*** (161.8)	611.1*** (161.9)	609.6*** (161.8)		107 (151)	126 (151)	141 (151)	
Debtdef		0.01 (15.53)	13.79 (14.96)	0.39 (15.23)	-15.26*** (4.081)	-14.44*** (4.094)	-15.53*** (4.077)		-9.08** (4.40)	-10.80** (4.42)	-10.53** (4.41)	
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	
Sample size	2535				10774				10973			

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimators.

4.2.2. 2001–2009 and 2010–2018

We then continue our analysis by dividing our sample into two time periods around the GFC. The results of this exercise can be found in [Table 4](#). In terms of the financial variables, the first notable difference from the main results is that the mean returns differentials are not statistically significant in the first period. The opposite is true for the correlations, which are significant in the first period but mostly no longer significant in the second period. Differences in variances remain significant in both periods.

In terms of macroeconomic variables, the *CA* share is significant in the first period but not in the second half after the GFC. Moreover, differences in GDP per capita are important in the first sub-sample but cease to be significant in the second period. Finally, the *BCS* is not significant in the first period but becomes significant in the second half of the period. This could be associated with the change in the significance of the correlation coefficient described above. In summary, we can see that there tends to be a mechanism that works against international risk sharing, confirming our previous results. In the case of the macroeconomic variables, we identify the flow of equities among countries in the same business cycle. In the case of institutional variables, there are no notable changes between the main results and when the sample is divided into two time periods.

4.3. Semi-parametric regression

To examine whether the results might be driven by nonlinearities, we turn to the results of the semi-parametric regression depicted in [Fig. 3](#). In the case of the three financial variables, placed in panels (a), (b), and (c), we can identify a straight line as the best nonlinear estimate. This, on the one hand, may validate the use of a linear estimator. On the other hand, we can also see that the confidence bands spread considerably as the observations move toward the lowest and highest values of the independent variables. This could mean that the shape of the line is mostly driven by medium size observations that dominate the sample. Consequently, the results show that using semi-parametric regression and a further examination of the results by quantiles may reveal some new facts about the underlying relationships between portfolio equity flows and financial variables.

The case for nonlinearities is even stronger among the macroeconomic variables. For instance, exports as a share of GDP grow almost linearly for the low values that dominate the sample, however, decreasing returns, and eventually a fall in the relationship is visible for the high bilateral export shares. In the case of the output gap differentials, the opposite is true. Only in the case of debt can we make a strong case for a linear estimator.

However, the most profound nonlinear effects are found in the instance of institutional variables. Capital controls are the least severe case with a visible convex shape for high values of the measure. In the case of reporting country stability and partner country voice and

Table 4
Results from sub-samples: 2001–2009 and 2010–2018.

Variable	2001–2009				2010–2018			
	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4
Rdif	2690 (2694)	2872 (2716)	1599 (2735)	2007 (2733)	11683** (4979)	10380** (4960)	11440** (4949)	9999** (4959)
SDdif	5865*** (1655)	5984*** (1751)	4439* (1741)	5406*** (1749)	13458*** (2742)	15970*** (2848)	13870*** (2838)	15730*** (2847)
Cor	1896*** (215)	780*** (233)	577* (240)	530** (240)	2244*** (275)	387 (306)	411 (310)	239 (308)
CA		−1663*** (436)	−1437*** (428)	−1428*** (428)		−1099* (608)	−323 (604)	−959 (608)
CapControls		1753*** (230)	1311*** (237)	1478*** (239)		1383*** (332)	1696*** (343)	1124*** (340)
Exchange		1941 (2117)	1489 (2119)	1496 (2117)		5674 (3881)	5869 (3903)	7028* (3897)
RepStab		159 (120)				576*** (135)		596*** (135)
ParStab		−374*** (115)		−764*** (143)		283** (130)		−149 (176)
RepVoice			292** (149)	261* (149)			−251 (174)	
RepVoice			243* (140)	803.1*** (175)			594*** (164)	(803)*** (222)
Export		31570*** (6207)	28730*** (6192)	32040*** (6216)		26830*** (7223)	26910*** (7221)	27870*** (7225)
Import		17700*** (5754)	18750*** (5744)	15020*** (5780)		26880*** (6713)	26240*** (6708)	23760*** (6765)
GDPpedif		−0.00*** (0.00)	−0.00*** (0.00)	−0.00*** (0.00)		−0.00 (0.00)	−0.00 (0.00)	−0.00 (0.00)
BCS		160 (139)	188 (140)	189 (140)		485*** (167)	493*** (167)	498*** (167)
Debtdef		−0.47 (3.87)	−1.99 (3.95)	−2.92 (3.95)		−20.18*** (4.75)	−14.96*** (4.79)	−18.34*** (4.77)
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)
Sample size	12141				12141			

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimator.

accountability, the results demonstrate positive associations, nevertheless, with a very high degree of irregularity. Consequently, we believe that examination of the results using quantile regression may prove to be illuminating in exploring these relationships.

4.4. Quantile regressions

We now report quantile regression results in Table 5. In the case of the financial variables, we see that differences in the lagged rates of returns are only important for the bottom and for the top quantiles, where the outflows are the highest, and the inflows are the highest, respectively. There is no statistical significance for the quantiles in between. For the lagged differences in standard deviations, we have statistically significant results in the bottom three quantiles, where all the data refers to capital outflows, and additionally at the ninth quantile, but only at the lowest conventional confidence level.

Nevertheless, the most interesting results are for the lagged correlations. The coefficient is positive and significant in the third, fourth, fifth, eighth, and ninth quantiles, and from the second to the ninth quantile, the coefficients are positive. This group is likely to be driving the sign of the coefficients in the main results. However, the situation is different in the bottom and top quantiles — the coefficient is negative and significant, providing evidence for risk-sharing behavior that could not be seen in the main results. What we can observe in these quantile results is that fundamental financial forces are not particularly

important for the medium size flows, however, they are crucial for the determination of the highest size of outflows and inflows.

A similar picture can be seen in the case of the current account as a share of GDP. The coefficients are only significant at the two bottom and two top quantiles. This points to the possibility that the CA is a relevant determinant for very big and very low capital flows.

Moreover, international business cycle theory predicts the existence of capital flows among countries in different phases of the business cycle. Again, we do not find this to be true in the middle quantiles, where there are capital flows among the countries in the same phase of the business cycle. However, at the bottom quantile and at the top two quantiles, the results are no longer statistically significant. Another interesting result is uncovered in the case of bilateral imports as a share of GDP. In the top and in the bottom two quantiles the higher the imports, the higher the purchases of the portfolio equity — in line with the predictions of macroeconomic fundamentals. However, the results for quantiles from fourth to ninth are positive, which may indicate the role of financial ties through trade.

Overall, we can see that the macroeconomic variables again tend to be important for the big inflows and outflows but not for what happens in the middle. Bilateral exchange rate is important in most of the quantiles, and always has a positive sign. This implies risk-loving behavior, but not at the lowest two quantiles. We also find interesting results for the differences in real GDP per capita. The flows from the richer to poorer countries, found in the main results, only occur in the case of bigger inflows — the results are not significant for the outflows

Table 5
Estimation results of the quantile regression.

Variable	1st quantile	2nd quantile	3rd quantile	4th quantile	5th quantile	6th quantile	7th quantile	8th quantile	9th quantile	10th quantile
Rdif	1488*** (405)	352 (253)	-342 (220)	-21.2 (167)	-82.0 (196)	-105 (360)	-345 (318)	-123 (325)	1166 (975)	54613** (23715)
SDdif	2319*** (402)	1055*** (244)	480*** (169)	62.5 (115)	-9.4 (173)	142 (309)	384 (337)	288 (215)	1320* (709)	12724 (9376)
Cor	-53.0* (30.4)	25.0 (17.9)	38.1** (15.9)	54.3*** (19.2)	58.7** (24.8)	12.9 (28.2)	18.5 (24.9)	135*** (30.1)	301*** (93.5)	-3917*** (639)
CA	-353*** (81)	-163*** (52)	-45.5 (38.9)	10.0 (34.8)	-16.5 (43.0)	33.4 (69.3)	64.2 (90.3)	131 (104)	343*** (120)	-4504*** (1719)
CapControls	454*** (56)	156*** (25)	111*** (24)	82.7*** (21)	120*** (21)	145*** (35)	120*** (30)	136*** (33)	71.1 (76.8)	7588*** (1304)
Exchange	-268 (246)	412* (235)	534*** (118)	362** (183)	541 (400)	987* (571)	820** (346)	657** (300)	-25.9 (574)	37715** (14824)
RepStab	-74.4*** (11.6)	-8.37 (9.62)	30.5*** (10.2)	57.3*** (8.9)	83.4*** (10.0)	108*** (14.8)	140*** (15.2)	128*** (17.6)	201*** (37.1)	1153 (1012)
ParVoice	27.3 (21.2)	18.0** (7.5)	37.1*** (10.1)	62.7*** (11.8)	101*** (17.0)	130*** (22.5)	166*** (25.8)	131*** (24.1)	158*** (49.5)	750 (934)
Export	-12671*** (4187)	-1594 (1274)	3481*** (1259)	7347*** (2001)	11580*** (2408)	18505*** (1812)	25952*** (2579)	39156*** (5208)	78161*** (10097)	-33136 (32536)
Import	-24712*** (3538)	-5184*** (1390)	-1481 (1186)	2349* (1304)	5412*** (1670)	12053*** (1640)	19746*** (3685)	42101*** (5094)	122325*** (13885)	-143330*** (24014)
GDPpcdif	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)	-0.00*** (0.00)	-0.00*** (0.00)	-0.00*** (0.00)	-0.00** (0.00)	-0.00** (0.00)
BCS	30.4 (18.8)	26.4** (12.7)	70.3*** (10.6)	74.6*** (10.7)	184*** (14.9)	203*** (19.1)	166*** (21.9)	84.8*** (20.1)	97.9 (60.9)	-704 (870)
DebtDif	-1.17 (0.72)	0.08 (0.46)	0.30 (0.32)	0.54 (0.36)	-0.001 (0.52)	-0.54 (0.54)	-0.26 (0.48)	-0.10 (0.49)	-2.52 (1.56)	-19.2 (23.5)
GDPprod	-0.00*** (0.00)	-0.00** (0.00)	0.00 (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	-0.00*** (0.00)
Sample size	2428	2428	2428	2428	2428	2428	2428	2428	2428	2428

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimators.

and small inflows – and the results become statistically significant only for the top 5 quantiles. This not only indicates that portfolio equity flows travel from the rich to poor countries but even more importantly, that they stay there. Finally, GDP product is significant and negative in the bottom two quantiles, and in the top one. Everywhere else, it is positive and significant, except for quantile three. This indicates that in the case of very big capital flows the size of the trading economies does not matter.

When it comes to institutional variables, we see that capital controls are always important (except for quantile nine) and always have a positive sign, as expected. The other two institutional variables are important only in the middle — where they are significant and with a positive sign. The only exception is a significant and negative coefficient on the reporter country stability, which should be treated as an anomaly. In summary, we have evidence that financial and overall macroeconomic variables are important determinants for big cross-country capital flows, while institutional variables are important determining factors for small capital flows.

5. Conclusions

In this paper, we have examined the data on bilateral inflows of portfolio equity among 40 developed economies over the period between 2001 and 2018. When we look at the entire sample, on the one hand, we observe financial inflows into countries with relatively higher returns, as predicted by the classical Markovitz model. On the other hand, we see, somewhat against the predictions of the model, that capital flows to countries with relatively more volatile returns. This result is at odds with the general notion of risk-averse economic agents and rather testifies to the risk-seeking behavior of the agents.

This outcome is not new to the literature (Crum et al., 1981), especially in the context of equity markets (Post and Levy, 2005).

The last prediction of the Markovitz model that economic agents will try to maintain assets characterized by low correlations in their portfolios is irreconcilable with our results. We report a positive role of the correlation on portfolio equity flows. On the one hand, this finding stands in contrast to the international business cycle literature (Backus et al., 1992; Backus and Smith, 1993), which underlines the role of risk-sharing by agents who diversify their portfolios internationally in order to achieve greater stability in their consumption path. On the other hand, this result provides the empirical grounds for the lack of consumption risk-sharing observed in macroeconomic data in the vast body of research (Kalemli-Ozcan et al., 2001; Afonso and Furceri, 2008; Kose et al., 2009; Dufrénot et al., 2020; Beck and Yersh, 2024).

Another result that is different to the predictions of international business cycle theory is the presence of intensified flows among countries within the same phase of the business cycle. Regardless of whether we approach this issue from the point of view of capital moving from places with depressed returns to economies with a higher yield, or from the perspective of *ex post* risk-sharing, for agents selling equity in the depressed countries and purchasing equity in countries experiencing an economic expansion, the movement of the capital should be observed across countries in different phases of the business cycle. However, the data shows otherwise, yet reinforcing the arguments against the presence of international consumption risk-sharing.

Turning to other macroeconomic factors, the influence of the country's position of the current account is in line with the economic theory, as countries with current account deficits attract higher capital flows. Similarly, close bilateral trade ties, whether proxied by exports or imports, contribute positively to the magnitude of cross-country portfolio equity inflows. On the other hand, the exchange rate variability

Table 6
Summary of the results per quantile.

Variable type			Financial			Institutional			Macroeconomic							
Quantile	From (mln)	To (mln)	Rdif	SDdif	Cor	CapControls	RepStab	ParVoice	CA	BCS	DebtDif	GDPpedif	Exchange	Export	Import	GDPprod
1st	-100738.600	-428.284	+	+	-	+	-		-	+				-	-	-
2nd	-428.284	-44.161		+		+		+					+		-	-
3rd	-44.161	-3.782		+	+	+	+	+		+			+	+		
4th	-3.782	0.003			+	+	+	+		+			+	+	+	+
5th	0.003	1.384			+	+	+	+		+				+	+	+
6th	1.384	16.752				+	+	+		+		-	+	+	+	+
7th	16.752	94.290				+	+	+		+		-	+	+	+	+
8th	94.290	446.400			+	+	+	+		+		-	+	+	+	+
9th	446.400	2421.793		+		+	+	+				-		+	+	+
10th	2421.793	169495.205	+		-	+						-	+		-	-
Total	-100738.600	169495.205	+	+	+	+	+	+		+	+	-		+	+	+

+ denotes positive and statistically significant coefficient, - denotes negative and statistically significant coefficient, while blank spaces represent coefficients not significant at any conventional level.

does not have an impact on the portfolio equity flows. We also report the significant influence of the differences in the level of economic development on the size of capital flows. Interestingly, in contrast to the (Lucas, 1990) paradox, we find that the movement of capital goes from the richer to the poorer countries.

Additionally, higher relative sovereign indebtedness of a country deters equity flows along the lines of standard economic theory predictions. Nevertheless, this result is strongly conditioned upon the sovereign rating level of the examined countries. The inflow of equity into the countries with the AAA rating is not affected by the difference in government debt ratios, as those countries are expected to make due on their obligations regardless of the size of their debt. However, countries with lower sovereign ratings must take into consideration their sovereign indebtedness when they want to attract additional capital, as the risk of insolvency might discourage potential investors.

Institutional factors also play an important role in driving portfolio equity flows. Capital controls still constitute one of the main forces behind the equity flows, but not in the countries with a AAA sovereign rating. This result is not surprising as the degree of capital mobility across those countries is very high with virtually no capital controls. A similar case can be made for the degree of political stability in the reporting country and the availability of reliable information in the partner country, which is essential in countries with a below AAA rating.

The nonlinearities we observed motivated us to examine the data within 10 quantiles. The summary of the results is depicted in Table 6. This exercise has proven to be extremely instructive, as many of the conclusions reached based on the full sample can be put into context. Moreover, as this is the first research that investigates the importance of determinants of capital flows conditioned on the magnitude of the flows, we are able to place more appropriate economic interpretations on the phenomena described in the preceding paragraphs. The main conclusion that can be taken from the results in quantiles is the difference between what drives the flows on the tails and what determines them in the middle of the distribution.

We report that the differences in the mean of returns and differences in their standard deviations are only important on the tails, while their role is insignificant in the middle of the sample. This outcome is most visible in the case of relative returns, which influence the flows only in the cases of the highest outflows and inflows. In the case of portfolio equity flows, we still observe the risk-seeking behavior of agents allocating resources in the countries with higher relative variances of the returns.

The results from the quantile regression shed a different light on the conclusions concerning the direction of capital flows and risk-sharing from the international business cycle models. Correlations of

the returns have a positive or no impact on the flows in the eight middle quantiles. However, in the very bottom and top quantiles, the coefficient turns negative, indicating the risk-sharing is in line with predictions of the standard models (Obstfeld and Rogoff, 1996). The capital flows across countries in the same phase of the business cycle are significantly higher for countries in the middle quantiles, while it is not the case in the bottom and in the two top quantiles. Accordingly, the predictions of international business cycle models work very well in the tails, ergo in the cases of outflows and inflows of the highest magnitude (approximately above 400 million Euros in absolute value). A similar statement can be made for the current account position that has a negative and significant impact on the portfolio equity flows in the two bottom and in the top quantile. Those results taken together show that the inflows and outflows of the highest magnitude are in fact influenced by major financial and macroeconomic forces along the lines of the prediction of the standard models, notwithstanding the risk-seeking behavior of the economic agents.

The results also show that the factors that are driving the medium-size flows are associated with institutions and bilateral relations across countries. For the medium quantiles, bilateral trade relations, proxied by imports and exports, positively influence the portfolio capital flows. The same can be inferred for exchange rate volatility, however, here we again find the risk-seeking behavior of the economic agents. The absence of capital controls, political stability in the reporting country, and availability of information in the partner country all have a positive impact on portfolio equity flows. Consequently, the role of institutional factors and bilateral relations is crucial in determining international portfolio equity flows.

Finally, we can make a very interesting observation about the role of differences in the degree of economic development across the examined economies. In contrast with the Lucas paradox, we see that capital flows from richer to poorer countries; however, this result is significant only in the top five quantiles. In other words, the difference in development among the economies matters only for inflows and not for outflows. Consequently, the portfolio capital flows from the richer countries, and once it is in the poorer countries, other factors determine the decision about its withdrawal.

Declaration of competing interest

The authors declare that they have no known competing financial interests or personal relationships that could have appeared to influence the work reported in this paper.

Data availability

Data will be made available on request.

Appendix A. Missing country pairs

Reporter	Partner	Reporter	Partner	Reporter	Partner	Reporter	Partner	Reporter	Partner
Australia	Estonia	Hong Kong	Hungary	Netherlands	Slovakia	Malta	Hungary	Turkey	Lithuania
Australia	Hungary	Hong Kong	Latvia	New Zealand	Norway	Mauritius	Hungary	Mexico	Malta
Australia	Latvia	Hong Kong	Norway	New Zealand	Panama	Mexico	Hungary	New Zealand	Malta
Australia	Lebanon	Hong Kong	Slovakia	New Zealand	Poland	Netherlands	Hungary	Romania	Malta
Australia	Norway	Hungary	Lebanon	New Zealand	Romania	New Zealand	Hungary	Singapore	Malta
Australia	Slovakia	Hungary	Norway	New Zealand	Slovakia	Norway	Hungary	Turkey	Malta
Austria	Hungary	Hungary	Slovakia	New Zealand	Turkey	Panama	Hungary	Mexico	Mauritius
Austria	Norway	Iceland	Latvia	Norway	Slovakia	Poland	Hungary	New Zealand	Mauritius
Austria	Slovakia	Iceland	Lebanon	Panama	Romania	Portugal	Hungary	Panama	Mauritius
Belgium	Hungary	Iceland	Malta	Panama	Slovakia	Romania	Hungary	Poland	Mauritius
Belgium	Norway	Iceland	Mauritius	Poland	Slovakia	Singapore	Hungary	Portugal	Mauritius
Belgium	Slovakia	Iceland	Norway	Portugal	Slovakia	Slovakia	Hungary	Romania	Mauritius
Canada	Hungary	Iceland	Romania	Romania	Slovakia	Spain	Hungary	Slovakia	Mauritius
Canada	Norway	Iceland	Slovakia	Singapore	Slovakia	Sweden	Hungary	Spain	Mauritius
Canada	Slovakia	Ireland	Norway	Malta	Chile	Switzerland	Hungary	Turkey	Mauritius
Chile	Hungary	Ireland	Slovakia	New Zealand	Chile	Turkey	Hungary	New Zealand	Mexico
Chile	Lebanon	Italy	Norway	Romania	Chile	UK	Hungary	Poland	New Zealand
Chile	Norway	Italy	Slovakia	Singapore	Chile	USA	Hungary	Romania	New Zealand
Chile	Slovakia	Japan	Norway	Mauritius	Czechia	New Zealand	Iceland	Panama	Norway
Czechia	Hungary	Japan	Slovakia	Mexico	Czechia	Singapore	Iceland	Poland	Norway
Czechia	Norway	Korea	Norway	New Zealand	Czechia	Turkey	Iceland	Portugal	Norway
Czechia	Slovakia	Korea	Slovakia	Panama	Czechia	Lebanon	Latvia	Romania	Norway
Estonia	Hungary	Latvia	Lebanon	Turkey	Czechia	Malta	Latvia	Singapore	Norway
Estonia	Lebanon	Latvia	Norway	Greece	Estonia	Mauritius	Latvia	Slovakia	Norway
Estonia	Norway	Latvia	Slovakia	Hong Kong	Estonia	Mexico	Latvia	Spain	Norway
Estonia	Slovakia	Lebanon	Norway	Lebanon	Estonia	New Zealand	Latvia	Sweden	Norway
Finland	Hungary	Lebanon	Slovakia	Mauritius	Estonia	Panama	Latvia	Switzerland	Norway
Finland	Lebanon	Lithuania	Norway	Mexico	Estonia	Romania	Latvia	Turkey	Norway
Finland	Norway	Lithuania	Slovakia	New Zealand	Estonia	Singapore	Latvia	UK	Norway
Finland	Slovakia	Luxembourg	Norway	Panama	Estonia	Turkey	Latvia	USA	Norway
France	Hungary	Luxembourg	Slovakia	Singapore	Estonia	Lithuania	Lebanon	Romania	Panama
France	Norway	Malta	New Zealand	Turkey	Estonia	Malta	Lebanon	Turkey	Panama
France	Slovakia	Malta	Norway	Malta	Hong Kong	Mexico	Lebanon	Singapore	Romania
Germany	Hungary	Malta	Panama	Romania	Hong Kong	New Zealand	Lebanon	Spain	Slovakia
Germany	Norway	Malta	Slovakia	Iceland	Hungary	Panama	Lebanon	Sweden	Slovakia
Germany	Slovakia	Mauritius	Norway	Ireland	Hungary	Poland	Lebanon	Switzerland	Slovakia
Greece	Hungary	Mauritius	Slovakia	Italy	Hungary	Portugal	Lebanon	Turkey	Slovakia
Greece	Lebanon	Mexico	New Zealand	Japan	Hungary	Romania	Lebanon	UK	Slovakia
Greece	Malta	Mexico	Norway	Korea	Hungary	Singapore	Lebanon	USA	Slovakia
Greece	Mauritius	Mexico	Poland	Latvia	Hungary	Slovakia	Lebanon		
Greece	Norway	Mexico	Romania	Lebanon	Hungary	Mauritius	Lithuania		
Greece	Panama	Mexico	Slovakia	Lithuania	Hungary	Mexico	Lithuania		
Greece	Slovakia	Netherlands	Norway	Luxembourg	Hungary	New Zealand	Lithuania		

Appendix B. List of countries and stock indices

Country	Index	Country	Index	Country	Index	Country	Index	Country	Index
Australia	AS51	France	CAC	Japan	NKY	Mexico	MEXBOL	Singapore	STI
Austria	ATX	Germany	DAX	Korea	KOSPI	Netherlands	AEX	Slovakia	SKSM
Belgium	BEL20	Greece	ASE	Latvia	RIGSE	New Zealand	NZSE	Spain	IBEX
Canada	SPTSX	Hong Kong	HSI	Lebanon	BLOM	Norway	OBX	Sweden	OMX
Chile	IGPA	Hungary	BUX	Lithuania	VILSE	Panama	BVPS	Switzerland	SMI
Czechia	PX	Iceland	ICEXI	Luxembourg	LUXXX	Poland	WIG	Turkey	XU100
Estonia	TALSE	Ireland	ISEQ	Malta	MALTEX	Portugal	PSI20	UK	UKX
Finland	HEX25	Italy	FTSEMIB	Mauritius	SEMDEX	Romania	BET	USA	SPX

Appendix C. Main results with stock market indices expressed in local currency

Variable	Model 1	Model 2	Model 3	Model 4
Rdif	5620** (2707)	6357** (2694)	4034** (2711)	4857* (2703)
SDdif	6521*** (1633)	7403*** (1707)	5108*** (1705)	6160*** (1707)
Cor	1770*** (164)	399** (176)	316* (178)	302* (177)
CA		-1285*** (367)	-770** (363)	-1129*** (368)
CapControls		1589*** (196)	1523*** (202)	1248*** (199)
Exchange		3096 (1978)	3198 (1978)	3160 (1977)
RepStab		369*** (90)		393*** (90)
ParStab		30.3 (86.5)		
RepVoice			-8.8 (113)	
ParVoice			484*** (106)	513*** (106)
Export		29330*** (4779)	28680*** (4776)	29060*** (4763)
Import		23780*** (4439)	23240*** (4440)	22010*** (4447)
GDPpcdif		-0.00*** (0.00)	-0.00*** (0.00)	-0.00*** (0.00)
BCS		342*** (109)	349*** (109)	354*** (109)
DebtDif		-9.53*** (2.99)	-9.13*** (2.99)	-9.40*** (2.99)
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)
Sample size			24282	

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimator.

Appendix D. Results from sub-samples: 2001–2009 and 2010–2018 with stock market indices expressed in local currency

Variable	2001–2009				2010–2018			
	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4
Rdif	468 (2876)	1134 (2936)	−945 (2975)	−214 (2975)	17905*** (5487)	16330** (5448)	16290*** (5447)	15740*** (5448)
SDdif	5780*** (1746)	5749*** (1842)	4015** (1847)	4838*** (1852)	6299* (3258)	10190*** (3478)	6799* (3469)	9514*** (3481)
Cor	1972*** (209)	908*** (226)	761*** (230)	734*** (230)	1573*** (250)	−18 (270)	−102 (271)	−119 (271)
CA		−1748*** (437)	−1533*** (430)	−1525*** (430)		−812 (610)	−25 (605)	−664 (611)
CapControls		1736*** (230)	1288*** (237)	1452*** (239)		1500*** (329)	1823*** (341)	1223*** (338)
Exchange		2020 (2116)	1547 (2119)	1554 (2117)		6105 (3885)	6454* (3905)	7470* (3900)
RepStab		151 (120)		−747*** (143)		565*** (135)		577*** (135)
ParStab		−360*** (115)				323** (130)		−118 (177)
RepVoice			270* (149)	239** (149)			−247 (173)	
ParVoice			258* (141)	800*** (175)			656*** (163)	817*** (222)
Export		31040*** (6205)	28120*** (6190)	31330*** (6214)		28120*** (7217)	28640*** (7219)	29030*** (7218)
Import		17520*** (5754)	18290*** (5745)	14610** (5782)		28290*** (6722)	27540*** (6720)	25010*** (6777)
GDPpcdif		−0.00*** (0.00)	−0.00*** (0.00)	−0.00*** (0.00)		−0.00 (0.00)	−0.00* (0.00)	−0.00 (0.00)
BCS		159 (140)	184 (140)	183 (140)		492*** (167)	504*** (167)	502*** (167)
DebtDif		−0.44 (3.88)	−1.86 (3.95)	−2.72 (3.95)		−18.39*** (4.92)	−12.47** (4.97)	−16.30*** (4.95)
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)
Sample size	12141				12141			

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimator.

Appendix E. Results from sub-samples: sovereign ratings with stock market indices expressed in local currency

Variable	Flows among AAA rating countries				Flows among below AAA rating				Flows among AAA and below AAA rating			
	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4
Rdif	11644 (16847)	15460 (16660)	15460 (15460)	15510 (16660)	7379** (3749)	7598** (3726)	6133 (3749)	8165*** (3713)	3465 (3809)	2936 (3811)	192 (3842)	883 (3825)
SDdif	20719** (9937)	19020** (9937)	19020** (9924)	19120** (9915)	2654 (2249)	3811 (2361)	2107 (2366)	4519* (2328)	10279*** (2320)	9471*** (2441)	7117*** (2435)	7998*** (2440)
Cor	5895*** (928)	2365** (1094)	2365** (1094)	2364** (1094)	1731*** (245)	576** (256)	485* (257)	573** (256)	1190*** (231)	113 (247)	-2 (249)	-15 (248)
CA		-1812 (1169)	-1812 (1154)	-1791 (1159)		1160* (672)	1837*** (669)	1114* (671)		-1956*** (477)	-1482*** (471)	-1721*** (478)
CapControls		2489** (1197)	2489 (2489)	2570* (1061)		1575*** (278)	1851*** (291)	1719*** (266)		1603*** (272)	1417*** (277)	1213*** (275)
Exchange		17080 (10550)	17080 (10550)	17120 (10550)		2195 (2771)	1994 (2775)	2108 (2771)		2795 (2798)	2865 (2794)	2637 (2794)
RepStab		83 (566)				576*** (127)		567*** (127)		263** (129)		314*** (128)
ParStab		-2144*** (492)		-2138*** (490)		221* (124)				167 (125)		
RepVoice			380 (450)				-53 (183)				21 (151)	
ParVoice			493 (419)				319* (169)				670*** (142)	710*** (142)
Export		78150*** (12150)	78150*** (78150)	77860*** (11990)		19150** (10180)	19620** (10190)	20320** (10160)		13530** (6147)	13580** (6141)	13730** (6124)
Import		21190* (11550)	21190*** (11550)	21310* (11530)		48850*** (9307)	50640*** (9315)	48710*** (9308)		11660** (5701)	9751* (5701)	9059 (5705)
GDPpcdif		-0.00 (0.00)	-0.00 (0.00)	-0.00 (0.00)		-0.00 (0.00)	-0.00* (0.00)	-0.00 (0.00)		-0.00*** (0.00)	-0.00*** (0.00)	-0.00*** (0.00)
BCS		84.9 (423)	84.9 (423)	87.2 (423)		625*** (162)	633*** (162)	628*** (162)		104 (151)	120 (151)	133 (151)
DebtDif		0.13 (15.49)	0.13 (15.34)	0.58 (15.18)		-14.62*** (4.15)	-13.55*** (4.17)	-15.06*** (4.14)		-8.58* (4.42)	-10.04** (4.43)	-9.74** (4.42)
GDPprod		0.00** (0.00)	0.00*** (0.00)	0.00 (0.00)		0.00*** (0.00)	0.00*** (0.00)	0.00 (0.00)		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)
Sample size		2535				10774				10973		

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimator.

Appendix F. Results from sub-samples: Investment grade

Subsample	Flows among Investment grade countries				Flows among below Investment grade				Flows among Investment and below Investment grade		
	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3	Model 4	Model 1	Model 2	Model 3
Rdif	9095** (3253)	9825*** (3226)	8037** (3236)	8472*** (3230)	12399 (17825)	-1723 (17500)	1164 (17450)	1837 (17460)	2706 (3685)	1902 (3655)	1294 (3664)
SDdif	14219*** (2176)	15110*** (2202)	13630*** (2197)	14230*** (2201)	-5716 (8994)	12640 (10840)	8578 (11310)	4382 (9983)	3793** (1832)	4080** (2053)	2613 (2049)
Cor	2119*** (219)	437* (242)	238 (247)	255 (244)	5266*** (1636)	1046 (1895)	1853 (1864)	1193 (1804)	1630*** (287)	980*** (301)	1211*** (303)
CA		-1487*** (424)	-1175*** (424)	-1282*** (425)		9801 (7041)	18930*** (6821)	17100*** (5395)		300 (753)	1489** (743)
CapControls		1430*** (249)	1158*** (254)	973*** (251)	4668*** (1716)	5586*** (1740)	4751*** (1652)		1683*** (314)	2252*** (320)	
Exchange		4664* (2544)	4707* (2542)	4514* (2542)	-4144 (14990)	-5877 (15070)	-3404 (15000)		367 (2911)	-569 (2917)	
RepStab		473*** (141)		465*** (141)	1354 (878)				445*** (134)		
ParStab		-121 (129)			-1142 (778)				387*** (129)		
RepVoice			87 (156)				-526 (1417)				-276 (185)
RepVoice			680*** (141)	674*** (141)			-1675 (1235)				157 (171)
Export		32930*** (5106)	31760*** (5095)	32540*** (5084)	77960 (113900)	78830 (114500)	84050 (114100)		-50260** (21450)	-47250** (21570)	
Import		22070*** (4781)	22040*** (4765)	21100*** (4768)	65290 (102600)	54550 (102200)	68930 (1013000)		66690*** (16100)	63490*** (16230)	
GDPpcdif		-0.00 (0.00)	-0.00* (0.00)	-0.00 (0.00)	-0.00** (0.00)	-0.00*** (0.00)	-0.00*** (0.00)		-0.00 (0.00)	-0.00 (0.00)	
BCS		258** (130)	278* (130)	285** (130)	1426 (1022)	1328 (1024)	1292 (1023)		497*** (191)	467** (191)	
DebtDif		-5.09 (3.95)	-4.70 (3.94)	-5.11 (3.94)	-109*** (28)	-93 *** (34)	-74 *** (22)		-12,95*** (4,28)	-12,57*** (4,28)	
GDPprod		0.00*** (0.00)	0.00*** (0.00)	0.00*** (0.00)	0.00 (0.00)	0.00 (0.00)	0.00 (0.00)		0.00 (0.00)	0.00 (0.00)	
Sample size		18530			5415				337		

Standard errors are in parentheses; */**/** denotes coefficient statistically significant at 0.9/0.09/0.99 level. All models were estimated with country-pair and time fixed effects estimator.

Appendix G. Summary statistics of the examined variables

Table 7
Summary statistics — part 1.

Country	Statistic	PEinflows	Rdif	SDdif	Cor	CapControls	RapVaA	RapStab	GDPpcdif	Exchange	GAPdif	DBdif
Australia	Max	19039	0.18	0.07	0.97	1.00	1.51	1.33	46822	0.34	1.87	41.71
	Mean	282	0.00	0.00	0.60	0.71	1.41	0.98	8773	0.06	0.00	-40.85
	SD	2875	0.02	0.03	0.30	0.20	0.05	0.12	21368	0.04	0.42	41.55
	Median	9	0.00	0.00	0.69	0.70	1.40	0.94	8717	0.05	0.01	-34.69
	Min	-40508	-0.06	-0.27	-0.68	0.12	1.35	0.86	-72686	0.01	-2.08	-199.48
Austria	Max	7733	0.16	0.11	0.98	1.00	1.48	1.36	48094	0.30	2.12	84.34
	Mean	129	0.00	0.00	0.61	0.89	1.39	1.14	11408	0.04	0.00	12.18
	SD	1101	0.02	0.03	0.28	0.21	0.04	0.16	21525	0.04	0.40	44.42
	Median	13	0.00	0.01	0.68	1.00	1.38	1.15	11979	0.04	0.01	21.81
	Min	-5994	-0.06	-0.22	-0.51	0.16	1.31	0.82	-72026	0.00	-1.79	-158.42
Belgium	Max	31975	0.18	0.08	0.98	1.00	1.48	1.26	43048	0.30	2.19	106.98
	Mean	387	0.00	-0.01	0.63	0.88	1.37	0.78	7045	0.04	0.00	39.09
	SD	3104	0.02	0.03	0.29	0.21	0.04	0.22	21546	0.04	0.42	44.36
	Median	4	0.00	0.00	0.72	1.00	1.36	0.80	8047	0.04	0.00	48.64
	Min	-14064	-0.07	-0.25	-0.64	0.14	1.32	0.40	-76802	0.00	-1.89	-132.61
Canada	Max	77924	0.19	0.07	0.96	1.00	1.67	1.27	47440	0.26	2.06	91.7
	Mean	606	0.00	-0.01	0.55	0.89	1.46	1.07	9119	0.05	0.00	18.04
	SD	5234	0.02	0.03	0.29	0.21	0.07	0.12	21789	0.03	0.45	44.43
	Median	15	0.00	-0.01	0.63	1.00	1.45	1.07	9047	0.05	0.01	27.30
	Min	-52518	-0.07	-0.27	-0.68	0.16	1.35	0.83	-75003	0.01	-1.96	-147.93
Chile	Max	14843	0.21	0.07	0.94	1.00	1.29	1.09	8045	0.26	2.22	25.76
	Mean	32	0.00	-0.01	0.48	0.74	1.07	0.55	-24637	0.06	0.00	-49.05
	SD	1551	0.02	0.03	0.29	0.20	0.09	0.19	21050	0.04	0.66	41.81
	Median	0	0.00	0.00	0.53	0.70	1.07	0.45	-23972	0.05	-0.05	-40.68
	Min	-14990	-0.07	-0.28	-0.55	0.12	0.89	0.33	-111460	0.01	-2.00	-218.51
Czechia	Max	44825	0.18	0.13	0.97	1.00	1.04	1.11	28394	0.28	2.09	43.90
	Mean	5056	0.00	0.01	0.57	0.87	0.97	0.95	-8536	0.06	0.00	-30.23
	SD	14136	0.02	0.03	0.29	0.21	0.08	0.18	21539	0.04	0.42	44.23
	Median	1	0.00	0.01	0.64	1.00	1.00	0.99	-8497	0.05	-0.01	-21.64
	Min	-4722	-0.05	-0.24	-0.56	0.13	0.76	0.33	-93585	0.01	-1.39	-200.40
Estonia	Max	44824	0.16	0.12	0.95	1.00	1.21	0.93	18657	0.27	2.43	10.56
	Mean	3059	0.01	0.00	0.48	0.90	1.11	0.69	-16601	0.05	0.00	-54.62
	SD	11280	0.02	0.03	0.27	0.21	0.06	0.11	21193	0.04	0.64	41.56
	Median	0	0.01	0.00	0.52	1.00	1.10	0.66	-16358	0.04	-0.09	-45.59
	Min	-455	-0.07	-0.22	-0.51	0.16	0.99	0.57	-100850	0.00	-1.73	-224.23
Finland	Max	8174	0.19	0.10	0.98	1.00	1.78	1.76	44959	0.30	1.48	63.58
	Mean	235	0.00	0.00	0.62	0.90	1.56	1.40	4781	0.04	0.00	-12.46
	SD	1136	0.02	0.03	0.29	0.21	0.08	0.24	21391	0.04	0.42	42.16
	Median	10	0.00	0.00	0.71	1.00	1.55	1.43	5884	0.04	-0.04	-6.50
	Min	-4628	-0.06	-0.28	-0.56	0.16	1.46	0.91	-77638	0.00	-1.23	-173.70
France	Max	61501	0.20	0.08	0.99	1.00	1.48	0.93	39649	0.30	2.22	98.08
	Mean	1427	0.00	0.00	0.63	0.89	1.23	0.43	2375	0.04	0.00	17.72
	SD	10242	0.02	0.03	0.32	0.21	0.10	0.26	21667	0.04	0.41	44.86
	Median	32	0.00	0.00	0.74	1.00	1.21	0.48	3341	0.04	0.01	25.72
	Min	-57580	-0.08	-0.29	-0.56	0.16	1.12	-0.10	-80554	0.00	-1.80	-138.64
Germany	Max	52191	0.20	0.10	0.98	1.00	1.50	1.41	43386	0.30	2.34	81.38
	Mean	2651	0.00	0.00	0.61	0.89	1.39	0.85	7503	0.04	0.00	6.77
	SD	8114	0.02	0.03	0.32	0.21	0.05	0.21	21531	0.04	0.44	44.87
	Median	183	0.00	0.00	0.72	1.00	1.40	0.85	8737	0.04	0.01	17.31
	Min	-25002	-0.09	-0.27	-0.49	0.16	1.30	0.58	-76464	0.00	-1.89	-171.19
Greece	Max	49346	0.15	0.12	0.96	1.00	1.19	0.88	27517	0.30	1.97	190.09
	Mean	81	-0.01	0.03	0.60	0.83	0.87	0.19	-13134	0.04	0.00	82.76
	SD	2618	0.02	0.04	0.25	0.22	0.18	0.38	22039	0.04	0.64	48.38
	Median	0.00	-0.01	0.03	0.67	1.00	0.89	0.05	-11843	0.04	-0.10	82.59
	Min	-16727	-0.08	-0.25	-0.38	0.12	0.62	-0.23	-99301	0.00	-1.61	-70.88
Hong Kong	Max	38088	0.19	0.06	0.94	1.00	0.73	1.34	58297	0.28	1.69	-2.63
	Mean	458	0.00	0.00	0.55	0.88	0.47	1.00	17259	0.04	0.00	-66.25
	SD	3167	0.02	0.03	0.28	0.21	0.16	0.15	21836	0.03	0.41	43.64
	Median	1	0.00	0.00	0.63	1.00	0.49	0.96	17305	0.04	0.00	-56.54
	Min	-17055	-0.08	-0.27	-0.72	0.16	-0.01	0.75	-61757	0.00	-1.67	-233.46
Hungary	Max	44822	0.18	0.12	0.98	1.00	1.18	1.26	14226	0.29	1.37	79.87
	Mean	6215	0.00	0.02	0.55	0.87	0.86	0.82	-17299	0.06	0.00	9.03
	SD	15459	0.02	0.03	0.28	0.21	0.26	0.18	21084	0.04	0.39	41.82
	Median	0.00	0.00	0.02	0.58	1.00	0.90	0.75	-16783	0.05	0.01	16.01
	Min	-806	-0.05	-0.22	-0.5	0.12	0.40	0.54	-105280	0.01	-1.75	-163.42

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Table 7 (continued).

Country	Statistic	PEinflows	Rdif	SDdif	Cor	CapControls	RapVaA	RapStab	GDPpcdif	Exchange	GAPdif	DBdif
Iceland	Max	1900	0.07	0.34	0.88	0.70	1.63	1.62	59879	0.34	2.67	137.6
	Mean	3	0.00	0.02	0.31	0.41	1.45	1.36	10318	0.07	0.00	35.43
	SD	250	0.05	0.07	0.30	0.26	0.07	0.15	21482	0.05	0.80	48.94
	Median	0.00	0.00	0.00	0.32	0.41	1.44	1.35	11650	0.06	-0.01	38.52
	Min	-2531	-0.23	-0.21	-0.64	0.07	1.36	1.02	-75660	0.01	-1.94	-169.33
Ireland	Max	103042	0.15	0.08	0.94	1.00	1.62	1.59	86276	0.30	2.34	119.52
	Mean	1952	0.00	-0.01	0.54	0.89	1.36	1.12	22791	0.04	0.00	-0.31
	SD	10855	0.02	0.03	0.30	0.21	0.09	0.20	24730	0.04	0.95	51.86
	Median	100	0.00	0.00	0.62	1.00	1.32	1.05	22965	0.04	0.23	4.74
	Min	-90419	-0.08	-0.27	-0.46	0.16	1.27	0.85	-69697	0.00	-2.30	-169.33
Italy	Max	60306	0.19	0.09	0.99	1.00	1.18	0.92	39203	0.30	2.15	135.3
	Mean	504	-0.01	0.01	0.61	0.89	1.02	0.48	919	0.04	0.00	57.63
	SD	7443	0.02	0.03	0.30	0.21	0.06	0.16	21924	0.04	0.41	43.66
	Median	8	-0.01	0.01	0.69	1.00	1.03	0.47	512	0.04	0.00	64.72
	Min	-88623	-0.07	-0.28	-0.48	0.16	0.91	0.27	-83937	0.00	-1.65	-99.60
Japan	Max	169495	0.22	0.05	0.95	1.00	1.11	1.20	39711	0.27	2.00	233.46
	Mean	1916	0.00	-0.01	0.43	0.89	1.02	1.03	1165	0.06	0.00	139.31
	SD	13123	0.02	0.03	0.35	0.21	0.05	0.08	22067	0.03	0.42	44.66
	Median	43	0.00	-0.01	0.47	1.00	1.01	1.02	921	0.05	0.01	139.60
	Min	-41117	-0.07	-0.3	-0.68	0.16	0.95	0.88	-85784	0.01	-1.92	-17.96
Korea	Max	33197	0.18	0.10	0.96	1.00	0.79	0.64	30925	0.27	1.57	41.16
	Mean	340	0.00	0.01	0.53	0.55	0.71	0.35	-4814	0.05	0.00	-32.91
	SD	2554	0.02	0.03	0.29	0.23	0.04	0.14	21557	0.04	0.42	44.24
	Median	7	0.00	0.01	0.59	0.48	0.72	0.39	-5372	0.05	0.02	-23.81
	Min	-24928	-0.06	-0.24	-0.62	0.07	0.63	0.11	-88925	0.01	-1.76	-193.82
Latvia	Max	1328	0.18	0.12	0.96	1.00	0.85	1.00	12285	0.28	2.33	45.90
	Mean	6	0.01	0.00	0.36	0.89	0.80	0.55	-22420	0.05	0.00	-32.58
	SD	108	0.02	0.03	0.32	0.21	0.04	0.21	21051	0.04	0.59	42.09
	Median	0.00	0.01	0.00	0.40	1.00	0.79	0.48	-21863	0.04	-0.11	-27.63
	Min	-1294	-0.06	-0.28	-0.66	0.15	0.70	0.20	-107690	0.00	-1.46	-195.37
Lebanon	Max	765	0.23	0.13	0.87	0.70	-0.28	-0.36	1624	0.28	2.3	181.87
	Mean	-1	0.00	-0.01	0.19	0.56	-0.43	-1.4	-32338	0.04	0.00	91.23
	SD	81	0.03	0.05	0.29	0.17	0.09	0.53	21414	0.03	0.92	47.82
	Median	0	0.00	-0.02	0.19	0.70	-0.42	-1.63	-31276	0.04	0.04	99.95
	Min	-750	-0.07	-0.28	-0.75	0.07	-0.66	-2.12	-119850	0.00	-2.84	-95.24

Explanations of the abbreviations are described in Section 3.1.

Table 8
Summary statistics — part 2.

Country	Statistic	PEinflows	Rdif	SDdif	Cor	CapControls	RapVaA	RapStab	GDPpcdif	Exchange	GAPdif	DBdif
Lithuania	Max	44823	0.16	0.12	0.95	1.00	1.00	1.05	17723	0.28	1.81	42.62
	Mean	7168	0.01	0.00	0.43	0.80	0.91	0.78	-18937	0.05	0.00	-31.85
	SD	16411	0.02	0.03	0.29	0.21	0.04	0.13	21261	0.04	0.50	41.72
	Median	0	0.01	0.00	0.47	0.82	0.90	0.77	-18555	0.04	-0.01	-26.17
	Min	-458	-0.07	-0.22	-0.42	0.12	0.85	0.42	-106110	0.00	-1.74	-198.80
Luxembourg	Max	71116	0.17	0.11	0.99	1.00	1.67	1.64	119850	0.30	2.20	21.92
	Mean	2381	0.00	0.01	0.60	0.89	1.56	1.42	71286	0.04	0.00	-48.97
	SD	9707	0.02	0.03	0.29	0.21	0.06	0.09	22620	0.04	0.44	43.73
	Median	334	0.00	0.01	0.70	1.00	1.55	1.41	72329	0.04	-0.01	-40.47
	Min	-64142	-0.07	-0.22	-0.58	0.16	1.41	1.32	19149	0.00	-1.88	-212.91
Malta	Max	13628	0.21	0.04	0.94	1.00	1.37	1.60	33022	0.29	1.75	65.35
	Mean	150	0.00	-0.02	0.34	0.73	1.20	1.26	-9748	0.05	0.00	-3.66
	SD	995	0.02	0.03	0.30	0.33	0.07	0.17	21635	0.04	0.66	44.25
	Median	2.00	0.00	-0.02	0.37	0.94	1.19	1.25	-10758	0.04	-0.05	4.84
	Min	-3013	-0.06	-0.31	-0.55	0.03	1.09	1.04	-96004	0.00	-2.21	-188.94
Mauritius	Max	9288	0.20	0.10	0.90	1.00	0.98	1.12	8564	0.26	2.84	66.17
	Mean	38	0.00	-0.02	0.29	0.74	0.85	0.90	-27600	0.05	0.00	-8.29
	SD	636	0.02	0.03	0.30	0.22	0.06	0.13	22038	0.03	0.64	45.41
	Median	0	0.00	-0.02	0.29	0.70	0.84	0.95	-27031	0.04	0.06	1.79
	Min	-3085	-0.06	-0.29	-0.62	0.12	0.73	0.64	-114290	0.00	-2.19	-172.91
Mexico	Max	20394	0.20	0.09	0.96	0.70	0.35	-0.06	4983	0.22	1.72	56.67
	Mean	30	0.00	0.01	0.56	0.62	0.13	-0.58	-33287	0.06	0.00	-24.37
	SD	1196	0.02	0.03	0.29	0.15	0.13	0.25	21083	0.04	0.43	44.92
	Median	0.00	0.00	0.01	0.64	0.70	0.13	-0.67	-31860	0.05	0.01	-19.18
	Min	-10682	-0.05	-0.26	-0.68	0.07	-0.08	-0.85	-114530	0.01	-2.12	-184.67

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Table 8 (continued).

Country	Statistic	PEinflows	Rdif	SDdif	Cor	CapControls	RapVaA	RapStab	GDPpcdif	Exchange	GAPdif	DBdif
Netherlands	Max	40044	0.18	0.09	0.99	1.00	1.70	1.76	54692	0.30	2.17	67.95
	Mean	855	0.00	0.00	0.64	0.89	1.54	1.04	14294	0.04	0.00	-7.50
	SD	4830	0.02	0.03	0.29	0.21	0.06	0.22	21621	0.04	0.41	44.47
	Median	55	0.00	0.00	0.74	1.00	1.54	0.97	15033	0.04	-0.01	1.44
	Min	-24468	-0.08	-0.25	-0.61	0.16	1.45	0.78	-68095	0.00	-1.43	-180.04
New Zealand	Max	16267	0.06	0.04	0.93	1.00	1.68	1.60	15409	0.25	1.55	35.19
	Mean	44	0.00	-0.01	0.60	0.97	1.56	1.36	-15116	0.05	0.00	-45.11
	SD	1402	0.01	0.02	0.24	0.11	0.06	0.15	18513	0.04	0.38	45.40
	Median	1	0.00	0.00	0.63	1.00	1.56	1.36	-11517	0.04	0.02	-42.62
	Min	-13419	-0.04	-0.11	-0.46	0.42	1.48	1.07	-88543	0.01	-1.4	-204.38
Norway	Max	22490	0.16	0.14	0.99	1.00	1.74	1.61	73720	0.27	2.13	51.37
	Mean	478	0.00	0.01	0.59	0.89	1.62	1.27	27072	0.06	0.00	-23.77
	SD	3932	0.02	0.03	0.28	0.21	0.08	0.13	22880	0.04	0.63	45.89
	Median	21	0.00	0.01	0.66	1.00	1.63	1.26	27486	0.05	0.00	-14.00
	Min	-57717	-0.07	-0.20	-0.5	0.16	1.49	1.12	-58751	0.01	-1.79	-203.67
Panama	Max	2242	0.23	0.02	0.69	1.00	0.63	0.41	11720	0.28	1.67	60.02
	Mean	15	0.00	-0.04	-0.01	0.89	0.54	0.09	-30062	0.04	0.00	-19.82
	SD	135	0.02	0.03	0.28	0.21	0.06	0.20	21149	0.03	0.64	47.01
	Median	0	0.00	-0.04	-0.02	1.00	0.54	0.07	-28906	0.04	0.16	-8.26
	Min	-639	-0.07	-0.34	-0.75	0.16	0.42	-0.18	-119510	0.00	-2.67	-197.22
Poland	Max	44810	0.17	0.13	0.98	0.70	1.11	1.07	13666	0.28	1.65	56.02
	Mean	4461	0.00	0.02	0.57	0.45	0.97	0.69	-21232	0.07	0.00	-12.60
	SD	13371	0.02	0.03	0.30	0.16	0.11	0.28	21301	0.04	0.47	43.18
	Median	1	0.00	0.02	0.64	0.45	1.01	0.71	-20647	0.06	-0.01	-4.72
	Min	-4673	-0.06	-0.24	-0.58	0.03	0.70	0.15	-109600	0.01	-1.93	-183.65
Portugal	Max	10890	0.18	0.08	0.97	1.00	1.46	1.44	24482	0.30	1.43	132.88
	Mean	158	-0.01	0.00	0.57	0.90	1.21	0.98	-12634	0.04	0.00	38.07
	SD	1096	0.02	0.03	0.29	0.21	0.13	0.22	21211	0.04	0.40	46.78
	Median	2.00	0.00	0.00	0.66	1.00	1.18	0.96	-12410	0.04	-0.02	40.20
	Min	-8646	-0.07	-0.27	-0.54	0.16	1.03	0.72	-95553	0.00	-1.56	-110.99
Romania	Max	44827	0.13	0.13	0.95	1.00	0.60	0.46	8879	0.22	1.77	29.80
	Mean	4438	0.01	0.02	0.50	0.79	0.44	0.15	-30321	0.06	0.00	-39.33
	SD	13375	0.03	0.04	0.30	0.29	0.08	0.17	20824	0.03	0.56	42.53
	Median	0	0.00	0.02	0.53	1.00	0.45	0.18	-29330	0.06	-0.14	-31.19
	Min	-223	-0.09	-0.20	-0.51	0.03	0.30	-0.38	-115810	0.01	-1.38	-196.00
Singapore	Max	38742	0.06	0.07	0.97	1.00	0.13	1.62	86647	0.27	2.07	109.40
	Mean	682	0.00	-0.01	0.58	0.91	-0.10	1.26	39103	0.05	0.00	35.90
	SD	4321	0.01	0.03	0.28	0.19	0.16	0.17	23990	0.03	0.71	42.25
	Median	43	0.00	0.00	0.64	1.00	-0.10	1.22	41212	0.05	0.01	44.14
	Min	-46169	-0.05	-0.18	-0.41	0.16	-0.39	0.88	-39125	0.01	-1.81	-135.78
Slovakia	Max	44816	0.23	0.10	0.90	0.75	0.98	1.12	18432	0.27	1.78	54.21
	Mean	734	0.00	0.00	0.33	0.56	0.93	0.91	-16262	0.05	0.00	-16.13
	SD	5591	0.03	0.03	0.35	0.22	0.04	0.15	20848	0.04	0.57	42.41
	Median	1	0.00	0.00	0.39	0.63	0.94	0.93	-15678	0.04	0.00	-7.36
	Min	-2032	-0.08	-0.27	-0.95	0.03	0.84	0.59	-101880	0.00	-1.80	-182.84
Spain	Max	31987	0.20	0.10	0.98	1.00	1.32	0.46	36622	0.30	1.67	100.64
	Mean	421	0.00	0.00	0.61	0.89	1.12	0.05	-4020	0.04	0.00	4.93
	SD	4073	0.02	0.03	0.28	0.21	0.10	0.28	21605	0.04	0.46	48.16
	Median	2.00	0.00	0.00	0.69	1.00	1.10	0.02	-4661	0.04	-0.05	10.72
	Min	-16716	-0.07	-0.27	-0.54	0.16	0.99	-0.47	-88357	0.00	-1.19	-149.24
Sweden	Max	44833	0.19	0.09	0.97	1.00	1.74	1.48	50023	0.24	2.33	49.84
	Mean	2128	0.00	0.00	0.61	0.89	1.58	1.18	11539	0.06	0.00	-20.33
	SD	8677	0.02	0.03	0.30	0.21	0.06	0.16	21531	0.04	0.42	45.68
	Median	36	0.00	0.00	0.70	1.00	1.57	1.15	12164	0.05	0.01	-10.29
	Min	-5438	-0.09	-0.26	-0.55	0.16	1.49	0.94	-69828	0.01	-1.72	-193.56
Switzerland	Max	32816	0.22	0.06	0.96	1.00	1.69	1.58	68923	0.28	2.42	55.65
	Mean	396	0.00	-0.02	0.58	0.89	1.55	1.32	28000	0.06	0.00	-17.14
	SD	3416	0.02	0.03	0.28	0.21	0.07	0.10	21879	0.03	0.45	46.36
	Median	31	0.00	-0.01	0.66	1.00	1.56	1.30	28709	0.05	-0.01	-6.95
	Min	-18608	-0.07	-0.32	-0.44	0.16	1.41	1.20	-52285	0.00	-1.57	-193.23
Turkey	Max	812	0.06	0.23	0.91	0.45	0.01	-0.59	13530	0.32	2.53	73.59
	Mean	-1	0.00	0.06	0.48	0.30	-0.24	-1.06	-27682	0.09	0.00	-25.92
	SD	82	0.03	0.05	0.27	0.14	0.25	0.38	21895	0.06	0.45	52.05
	Median	0	0.00	0.05	0.49	0.31	-0.17	-0.94	-27291	0.07	0.00	-15.63
	Min	-512	-0.09	-0.08	-0.38	0.03	-0.85	-2.01	-113180	0.01	-1.83	-205.04

(continued on next page)

Table 8 (continued).

Country	Statistic	PEinflows	Rdif	SDdif	Cor	CapControls	RapVaA	RapStab	GDPpcdif	Exchange	GAPdif	DBdif
UK	Max	103526	0.19	0.05	0.98	1.00	1.60	1.04	41709	0.32	2.08	85.96
	Mean	3083	0.00	-0.02	0.63	0.89	1.34	0.42	4093	0.05	0.00	-0.40
	SD	13727	0.02	0.03	0.29	0.21	0.07	0.23	21698	0.04	0.43	46.55
	Median	164	0.00	-0.01	0.72	1.00	1.33	0.41	5318	0.04	0.03	5.67
	Min	-100739	-0.07	-0.31	-0.52	0.16	1.28	0.08	-81022	0.01	-1.88	-148.07
USA	Max	153978	0.21	0.05	0.96	1.00	1.34	1.08	62979	0.28	2.05	107.42
	Mean	3819	0.00	-0.02	0.56	0.89	1.16	0.43	22478	0.04	0.00	22.41
	SD	16827	0.02	0.03	0.33	0.21	0.11	0.29	21478	0.03	0.44	46.09
	Median	292	0.00	-0.02	0.67	1.00	1.12	0.44	22768	0.04	0.03	28.28
	Min	-71711	-0.08	-0.31	-0.48	0.16	0.98	-0.23	-62901	0.00	-2.02	-128.96

Explanations of the abbreviations are described in Section 3.1.

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